INCENTIVIZED TORTS: AN EMPIRICAL ANALYSIS

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ABSTRACT—Courts and scholars assume that group causation theories deter wrongdoers. This Article empirically tests, and rejects, this assumption, using a series of incentivized laboratory experiments. Contrary to common belief and theory, data from over 200 subjects show that group liability can encourage tortious behavior and incentivize individuals to act with as many tortfeasors as possible. We find that subjects can be just as likely to commit a tort under a liability regime as they would be when facing no tort liability. Group liability can also incentivize a tort by making subjects perceive it as fairer to victims and society. These findings are consistent across a series of robustness checks, including both regression analyses and nonparametric tests.

We also test courts’ and scholars’ insistence that the but-for test fails in cases subject to group causation. We use a novel experimental design that allows us to test whether, and to what extent, each individual’s decision to engage in a tortious activity is influenced by the decisions of others. Upending conventional belief, we find strong evidence that the but-for test operates in group causation settings (e.g., concurrent causes). Moreover, across our experiments, subjects’ reliance on but-for causation produced the very tort that group liability attempted to discourage.

A major function of liability in torts, criminal law, and other areas of the law is to deter actors from engaging in socially undesirable activities. The same is said about doctrines that result in group liability. Our empirical results challenge this basic logic.

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& Economics and Behavioral Economics Workshops, the American, European, Canadian, and Midwestern Law and Economics conferences, the 5th International Meeting in Law & Economics, Queen’s Law Faculty Works in Progress Seminar, and the Southern Economic Association Conference for their comments; and Shannon McLaughlin, Joshua Polk, Jennifer Sandlin, Laura Stephenson, and Stephen Walker for their research assistance.

INTRODUCTION

This Article undertakes to test two facially appealing, though empirically unsupported, assertions that courts and scholars accept and rely upon.

The first assertion is that group liability deters tortious behavior. Courts impose such liability using theories such as concerted action, substantial factor, and alternative liability.¹ The second assertion is that in cases subject

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¹ Under concerted-action theory, a number of parties can be held liable if they acted tortiously pursuant to an agreement “or to accomplish a particular result.” RESTATEMENT (SECOND) OF TORTS § 876 cmt. a (AM. L. INST. 1979). Even absent an agreement, a group of careless actors may be held liable under the substantial-factor doctrine if their independent careless behaviors concurred and each substantially contributed to the victim’s harm. See RESTATEMENT (SECOND) OF TORTS § 432(2) (AM. L. INST. 1965). Alternative liability allows courts to hold liable careless parties who did not injure the victim if the identity of the injurer is unknown. It does so by shifting the burden of proof to each actor to establish that she did not cause the victim’s harm. Id. § 433B(3); see also infra notes 4–7, 18–20 and accompanying text (discussing group-liability doctrines).
to group causation theories, the but-for test is inapplicable.\textsuperscript{2} For example, based on a belief in its ability to discourage antisocial behavior,\textsuperscript{3} courts hold liable all participants to a drag race, including non-injuring drivers and spectators.\textsuperscript{4} Similarly, courts have imposed liability on all actors who carelessly set fires, although each fire alone could have destroyed the victim’s property.\textsuperscript{5} They have also imposed liability on a group of hunters who independently shot in the plaintiff’s direction, although it was clear that only one of them injured the victim.\textsuperscript{6} The logic of these cases extends beyond their particular facts to many settings involving multiple actors.\textsuperscript{7}

\textsuperscript{2} See Boim v. Holy Land Found. for Relief & Dev., 549 F.3d 685, 695–97 (7th Cir. 2008) (Posner, J.) (explaining that in concerted-action, concurrent-forces, and alternative-liability cases, “the requirement of proving [actual] causation is relaxed”); see also infra notes 3–11.

\textsuperscript{3} The leading example is \textit{Summers v. Tice}, 199 P.2d 1, 1–3 (Cal. 1948), which held jointly and severally liable two hunters who carelessly shot in the direction of the plaintiff although it was clear that only one of them hit him. See DAN B. DOBBS, PAUL T. HAYDEN & ELLEN M. BUBLICK, DOBBS’ LAW OF TORTS § 193 (2d ed. 2020) (“The leading case in the United States is Summers v. Tice . . . .”); David A. Fischer, \textit{Products Liability—An Analysis of Market Share Liability}, 34 \textit{VAND. L. REV.} 1623, 1629, 1633 (1981) (noting that, in cases like \textit{Summers}, “each defendant’s conduct was unquestionably of an antisocial nature” and arguing that “the law seeks to . . . discourage socially undesirable activity”).

\textsuperscript{4} See \textit{RESTATEMENT (SECOND) OF TORTS} § 876 cmt. a, illus. 2 & cmt. d, illus. 4 (AM. L. INST. 1979); \textit{RESTATEMENT (THIRD) OF TORTS: APPORTIONMENT OF LIAB.} § 15 cmt. a, illus. 1 (AM. L. INST. 2000); See DOBBS ET AL., supra note 3, § 435 (explaining that in illegal drag races, all drivers “are jointly and severally liable to the plaintiff, although only one of them actually made contact with the plaintiff’s car”); see also Abel v. Eli Lilly & Co., 343 N.W.2d 164, 176 n.19 (Mich. 1984) (explaining that concerted-action theory “seems to have developed to deter hazardous group behavior”), superseded by statute, 1995 Mich. Pub. Acts 97, 102–03 (codified at Mich. COMP. LAWS § 600.6304 (1996)).

\textsuperscript{5} See DOBBS ET AL., supra note 3, § 189 (“In the example of the two fires that combine to burn the plaintiff’s property, the substantial factor test allows courts to avoid but-for analysis and to hold that the two tortfeasors who set the two different fires are both causes of the plaintiff’s harm, provided only that each fire was sufficient standing alone to cause the same harm.”); Anderson v. Minneapolis, St. P. & S.S.M. Ry. Co., 179 N.W. 45, 49 (Minn. 1920); \textit{RESTATEMENT (SECOND) OF TORTS} § 432(2) (AM. L. INST. 1965) (stating the substantial-factor test); \textit{RESTATEMENT (THIRD) OF TORTS: LIAB. FOR PHYSICAL & EMOTIONAL HARM} § 27 cmt. d (AM. L. INST. 2010) (“Imposing liability in such a circumstance serves the policy of deterrence.”); DOBBS ET AL., supra note 3, § 189 (referencing to the two-fire hypothetical as the “classic example” of a case in which but-for causation fails and courts substitute the substantial-factor test); W. PAGE KEETON, DAN B. DOBBS, ROBERT E. KEETON & DAVID G. OWEN, PROSSER AND KEETON ON THE LAW OF TORTS § 42 (5th ed. 1984) (using Anderson as an example of the substantial-factor test); Landers v. E. Tex. Salt Water Disposal Co., 248 S.W.2d 731, 735 (Tex. 1952) (applying the substantial-factor test in a case of multiple polluters).

\textsuperscript{6} \textit{Summers}, 199 P.2d at 1–3; see also DOBBS ET AL., supra note 4, § 193 (referring to \textit{Summers} as the “leading case in the United States” on alternative causation); Steven Shavell, \textit{An Analysis of Causation and the Scope of Liability in the Law of Torts}, 9 J. LEGAL STUD. 463, 494 (1980) (justifying the imposition of group liability in \textit{Summers} and noting that “[t]o do otherwise would inappropriately weaken the incentive of injurers to avoid harm”); Fischer, supra note 3, at 1629, 1633 (noting the role of the law in discouraging certain behaviors and arguing that \textit{Summers} is a “perfect example” of a case involving moral blame for antisocial conduct).

\textsuperscript{7} See, e.g., DAN B. DOBBS, \textit{THE LAW OF TORTS} § 175 (2000) (discussing the alternative-liability doctrine and noting that “the practical effect of the rule is to impose . . . liability upon each defendant in
Leading scholars also justify the application of group liability on theoretical grounds. For example, Professor Steven Shavell and Judge Richard Posner each explain that, absent group causation theories like alternative liability and concerted action, “the incentive of injurers to avoid harm[ing] others would be ‘inappropriately weaken[ed].’” Similarly, Professors Mark Grady, Saul Levmore, and Daryl Levinson each laud the ability of “collective sanctions” to deter group members from engaging in wrongdoing.

This Article first undertakes to test the theoretical assertion that group liability deters actors. Courts and scholars accept and rely on this assertion—perhaps because of its intuitive appeal—despite the lack of empirical support. The Article also tests the claim that the but-for test fails when group liability applies. Courts and scholars alike insist that “a literal and simple version of the but-for test” will exempt from liability the spectators in the drag race, the actors who carelessly set the fires, and the hunters in Summers v. Tice. The claim is that none of these actors are necessary and thus none can be a but-for cause of the harm.

Despite their widespread acceptance, the two claims—that group liability deters and that the but-for test fails in such settings—have never been tested. To date, they are based on conjecture and belief. This Article empirically tests these claims, using a novel series of incentivized laboratory experiments.

The numerous cases in which one defendant cannot show that the other defendant was the cause” (emphasis added)).

8 Shavell, supra note 6, at 494; RICHARD A. POSNER, ECONOMIC ANALYSIS OF LAW 212 (9th ed. 2014) (explaining that deterrence would be reduced without such theories); STEVEN SHAVELL, ECONOMIC ANALYSIS OF ACCIDENT LAW 108, 164–67 (1987) (discussing the economics of causation and the use of liability to induce “optimal behavior”).

9 Mark F. Grady, Res Ipsa Loquitur and Compliance Error, 142 U. PA. L. REV. 887, 913 (1994) (explaining that “defendants would have too little incentive to use precaution” without a theory such as res ipsa loquitur); Saul Levmore, Gomorrah to Ybarra and More: Overextraction and the Puzzle of Immoderate Group Liability, 81 VA. L. REV. 1561, 1563–64 (1995) (discussing res ipsa and justifying overextraction that may enhance deterrence); Daryl J. Levinson, Collective Sanctions, 56 STAN. L. REV. 345, 350 (2003) (justifying the application of collective sanctions on the grounds that they can “effectively deter wrongdoing”).

10 DOBBS ET AL., supra note 3, § 189 (taking the traditional view that “[w]hen each of two or more causes would be sufficient, standing alone, to cause the plaintiff’s harm, a literal and simple version of the but-for test holds that neither defendant’s act is a cause of the harm” and noting that “[t]he classic example is the case of two fires being swept by winds towards the plaintiff’s property”); see supra note 6 and accompanying text; KEETON ET AL., supra note 5, § 42.

11 In other words, the claim is that the injury would have occurred regardless of such actors. See, e.g., Boim v. Holy Land Found. for Relief & Dev., 549 F.3d 685, 695–97 (7th Cir. 2008) (explaining that “[i]n all these cases the requirement of proving [actual] causation is relaxed because otherwise there would be a wrong and an injury but no remedy”).
experiments. Similar to other concurrent-cause cases, in our experiments, each subject alone could cause the victim’s harm. To ensure the robustness of our results, we used two framings: a “Context-Free,” or “neutral,” frame and a contextualized variant, or a “Vignette,” involving polluting factories. In each framing, four out of every five subjects chose independently and simultaneously between a harmless activity with low payoffs and a tortious activity that promised higher gains. The payoffs of the fifth subject, the victim, were reduced to zero if at least one of the other four subjects chose the tortious activity.

Those who chose the tortious activity were subject to one of two liability regimes. In the No Liability regime, behaving tortiously came at no (monetary) cost to the tortfeasor. Accordingly, each subject had a dominant strategy to behave tortiously. Under the Liability regime, the tortfeasors had to fully compensate the victim in equal shares. Because the expected liability of each tortfeasor was diluted as more joined the activity, the payoffs depended on the decisions of others. Specifically, acting tortiously alone or with one other tortfeasor promised a loss. By contrast, when acting together with two or more tortfeasors, the expected liability of each was diluted to the point each could expect a gain. Importantly, subjects had to decide whether to join the tortious activity without knowing what others would choose.

Our novel experimental design allowed us to observe each subject’s preferences. We did so by employing the Strategy Method. This method allowed us to observe whether a subject’s willingness to act tortiously was conditional upon others doing the same. For example, we could observe whether a subject was willing to act only if two other actors joined her. The method also allowed us to check whether, despite not knowing others’ decisions, she was nevertheless willing to take a leap of faith and act tortiously.

Contrary to common belief, data from over 200 subjects show that group liability can encourage tortious behaviors. In our experiments, the majority of subjects preferred to act tortiously, but only if they were joined by others. Moreover, their willingness to engage in the tortious activities increased with the number of tortfeasors. In the neutrally framed experiment with liability, very few subjects were willing to act alone or with only one

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12 See supra note 5 and accompanying text.
13 This contextualized Vignette allowed us to test for the effects of group liability in a more realistic setting. The Vignette is described in Example 1 below. See infra Section I.A; see also infra Section II.B (discussing the experimental design); infra notes 80–81 and accompanying text (explaining the difference in the treatments’ contextual framing and potential changes in behavior elicited by contextualized instructions in economics experiments).
14 See infra Section II.B.4 (explaining the use of the Strategy Method to elicit subjects’ conditional and unconditional preferences).
additional actor. However, conditional upon two other subjects joining the tortious activity—when the activity became profitable—the proportions of subjects willing to act increased to 92%. Conditional on three others acting, it reached 96%. In other words, despite the imposition of liability, almost everyone was willing to engage in the tortious activity provided a large enough number of actors would do the same. The result in the contextualized framing (i.e., the polluting-factories setting) followed the same pattern. Moreover, we found that in certain situations, subjects were more willing to act tortiously under a Liability regime compared to a No Liability regime. The proportion of subjects willing to act conditional on two or three other subjects acting was higher under the Liability regime compared to the No Liability regime.

Subjects’ unconditional decisions were also surprising. Although most subjects preferred to act tortiously only as part of a group, the vast majority decided to behave tortiously unconditionally. They did so despite the fact that they could not communicate with each other. This suggests that subjects not only were encouraged to behave tortiously, but they also acted as if they believed that others would also be enticed by the imposition of group liability.

Further, upending conventional belief, the data also provide overwhelming evidence supporting the claim that the but-for test can be operable in group causation theories like substantial factor. In all of the liability sessions, regardless of framing, many subjects were unwilling to engage in the tortious activity but for at least two other subjects doing the same.

In our experiments, group liability seemed to impact subjects’ cost–benefit calculus in a number of ways. First, as the number of subjects increased, the amount each could expect to pay decreased because liability was shared by all. Consequently, the financial gain from the tortious activity increased. In other words, with more tortfeasors, each tortfeasor stood to profit more. Second, not only did individual profits increase with the number of tortfeasors, social gains could have increased in the same way. Thus, subjects may justify the tortious activity (e.g., pollution due to production of cement) as a necessary evil: the sacrifice of one (who is fully compensated) for the benefit of many. Group liability may have also reduced the psychological—that is, nonmonetary—cost associated with the tortious activity. Under the Liability regime, the victim was compensated, which may have reduced actors’ moral hesitation or psychic cost from harming another. Additionally, we observed that in the Liability regime, subjects’ reported fairness levels increased monotonically with the number of actors. In other
words, the more tortfeasors joined the activity, the more just it was considered: to the actors, the victim, and society.

This Article contributes to the economic and empirical research on tort law in a number of important ways. First, to our knowledge, this Article is the first to test the effect of group causation theories on actors’ behavior in a novel experimental setting. Previous studies have focused on the effect of tort law generally on actors’ decisions to engage in tortious activities but, due to design limitations, could not focus directly on the effect of causation law.  

Second, with few exceptions, earlier studies focused on a single tortfeasor. They could not therefore analyze the effect of causation law on torts involving multiple actors. Third, our study employs a unique design that allows us to observe actors’ causal decisions at the individual level. Fourth, unlike previous studies, we also test whether and to what extent an individual’s decision to engage in a tortious activity is influenced by the decisions of others.

The remainder of this Article proceeds as follows. Part I discusses the theory and is divided into two Sections. The first focuses on the conditions and mechanisms under which liability may encourage the commission of torts with multiple actors. The second centers on the actual-causation requirement and the applicability of the but-for test in cases involving multiple actors. Part II reviews the relevant empirical literature, describes the experimental design, presents the testable hypotheses and the theoretical predictions, and summarizes the results. The Article then provides concluding remarks. Our findings raise concerns about the desirability of group liability as a deterrent and call for caution in its application. Group liability may increase both the willingness to engage in the tortious activity and the level of perceived fairness to the actors, the victim, and society. Our findings also suggest that the but-for test can function, and even play an important role, in cases subject to group liability. Indeed, in our experiments, subjects were not willing to engage in the tortious activity but for others joining them, suggesting that each actor may have been the marginal but-for cause of the victim’s harm.

15 Prior literature and its limitations are discussed in Section II.A.
16 Contra Restatement (Third) of Torts: Liability for Physical & Emotional Harm § 27 reporters’ note, cmt. a (Am. L. Inst. 2010) (“There is near-universal recognition of the inappropriateness of the but-for standard for factual causation when multiple sufficient causes exist.”).
I. THE THEORY

A. Group Liability and Incentivized Torts

A major insight of the economic account of tort law is its role as a deterrent. By imposing appropriate costs on injurers, tort law incentivizes them to take care and avoid harming others. Economic accounts of tort law also contemplate that actors with private information may decide to commit torts when the benefits exceed the costs of liability. Courts and commentators have generally focused on the deterrence rationale in extending group liability, while largely ignoring the possibility that tort law can incentivize rather than deter some tortious activity. We use a stylized example below to illustrate how this second consequence from tort law can arise and demonstrate its link to causation in group torts.

In some cases, imposing liability on a group of wrongdoers can produce a dilution effect that can incentivize actors to join in committing a wrongdoing—a “tortfest” dynamic. As illustrated in our example below, in

17 See, e.g., POSNER, supra note 8, at 212 (“[I]t would be an economic mistake to let both [negligent parties] off scot-free, as that would reduce deterrence.”); MICHAEL FAURE, TORT LAW AND ECONOMICS, at xxiv–xxvi, 84–85 (2009) (discussing the historical development of economic analysis of tort law and the focus on “economic efficiency of preventive measures”); SHAVELL, supra note 8, at 5–32, 108 (considering how the scope of liability affects the behavior of parties); WILLIAM M. LANDES & RICHARD A. POSNER, THE ECONOMIC STRUCTURE OF TORT LAW 229 (1987) (“If the purpose of tort law is to promote economic efficiency, a defendant’s conduct will be deemed the cause of an injury when making him liable for the consequences of the injury will promote an efficient allocation of resources to safety and care.”); William M. Landes & Richard A. Posner, Joint and Multiple Tortfeasors: An Economic Analysis, 9 J. LEGAL STUD. 517, 521 (1980) (“[I]n the absence of any liability rules victims would bear the full costs of the accident.”).

18 See GUIDO CALABRESI, THE COSTS OF ACCIDENTS: A LEGAL AND ECONOMIC ANALYSIS 26–27, 68 (1970) (explaining that “the principal function of accident law is to reduce the sum of the costs of accidents and the costs of avoiding accidents” and noting that “the primary way in which a society may seek to reduce accident costs is to discourage activities that are ‘accident prone’ and substitute safer activities as well as safer ways of engaging in the same activities”); SHAVELL, supra note 8, at 108, 164–67 (discussing how liability in tort law incentivizes due care in cases where multiple injurers are involved); Richard A. Posner, A Theory of Negligence, 1 J. LEGAL STUD. 29, 32–33 (1972) (arguing that “the dominant function of the fault system is to generate rules of liability that if followed will bring about, at least approximately, the efficient—the cost-justified—level of accidents and safety”).

19 See, e.g., A. MITCHELL POLINSKY, AN INTRODUCTION TO LAW AND ECONOMICS 17–27, 43–51 (3d ed. 2003).


In some cases, group liability may not only lead to the tort being profitable to the tortfeasors, it may also be welfare-enhancing, and more so when the social value from committing the tort increases with the number of tortfeasors.

**Example 1.** Several actors (e.g., Factories) must choose between two activities: (1) a harmless activity (e.g., producing limestone) or (2) an alternative tortious activity (e.g., producing cement—a necessary ingredient for construction—that causes pollution). The harmless activity promises a $24 benefit for each actor, whereas the alternative activity promises $35. However, if one or more actors engage in the alternative activity, an innocent victim will incur an expected harm of $24. The alternative activity can trigger liability, shared equally by all participants in the tortious activity.\(^\text{22}\)

Standard rational-choice theory predicts that if only one or two actors are expected to choose the tortious activity, actors will engage in the harmless activity. In the case of one actor, the decision whether to engage in the harmless activity or the alternative activity is easy. Engaging in the harmless activity promises an expected gain of $24. By contrast, the alternative activity promises only $11, the difference between the $35 expected benefit and the $24 expected liability.

With two actors, the incentive to engage in the harmless activity remains strong. If both parties pursue the alternative activity, both will be held liable using a group causation theory. For example, in the case of polluting factories, courts may apply the substantial-factor doctrine.\(^\text{23}\) If the actors are participants (drivers or spectators) in a drag race during which one of the drivers injured the victim, both would be considered as acting in concert.\(^\text{24}\) And in the case of careless hunters where only one injured the victim, both would be liable under a theory of alternative liability.\(^\text{25}\) As a result, each actor can expect to pay half of the victim’s damages, $12 (24/2), and therefore receive a gain of $23 (35 – 12).\(^\text{26}\) They would thus prefer the harmless activity (24 > 23).

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\(^\text{22}\) For simplicity, this example assumes that taking care is impossible or too costly.

\(^\text{23}\) See *supra* note 5 and accompanying text.

\(^\text{24}\) See *supra* note 4 and accompanying text.

\(^\text{25}\) See *supra* note 6 and accompanying text.

\(^\text{26}\) With some simplifying assumptions, such as solvency, it is easy to show that the result is independent of the apportionment regime (e.g., several liability or joint and several liability, with or
However, once three actors are expected to participate, engaging in the alternative and harmful activity becomes a winning strategy: it promises each actor more profit than the harmless activity. If all three engage in the alternative activity, all would be held liable with the aid of a group causation theory. As a result, the expected liability of each will be diluted to $8 (24/3). This means that each of the three tortfeasors can expect to gain $27 (35 – 8)—more than the $24 gain promised by the harmless activity. Our example illustrates another point: the tortious activity can become more profitable and enticing as more actors join the activity if the expected liability decreases with the number of actors.

B. Actual Causation

Example 1 also illustrates the challenge to the consensus that, in group-liability settings (e.g., concerted action), the actual-causation requirement is missing. In fact, the example proves the consensus wrong. First, consider cases involving concurring causes like polluting factories—the subject of our Vignette. The traditional argument is that no one factory can be a but-for reason for the victim’s harm because the harm would have materialized in

27 For example, with four actors, each can expect to gain $29 (35 – (24/4)); with ten actors, the expected gain increases to $32.60 (35 – (24/10)). Note that in Example 1, the tortfeasor’s and society’s interests are aligned. With one or two participants, the tort is welfare decreasing. It would generate at most a net benefit of $46—the difference between the $70 (35 • 2) benefit to the tortfeasors and the $24 cost to the victim. By contrast, with three (or more) actors, the tortious activity becomes socially desirable. It would generate a net benefit of $81 ((35 • 3) – 24) compared to a net gain of only $72 (24 • 3) from the harmless activity. Of course, the social acceptability of a tort may change with the nature of the activity. It is possible that even if many agree that a harmful activity, like producing cement, is justified, fewer would agree that the same is true for hunting, and fewer, if any, would defend injurious drag races in the name of efficiency.

28 The model assumes that the parties know that the tortious activity (e.g., pollution) is subject to liability and that they are able to engage in an ex ante cost–benefit analysis. Importantly, the model does not assume that parties (or courts) are omniscient. Rather, it relies on the fact that they are not. Each actor is assumed to know her expected benefits from the activity (but not others’ benefits or decisions) and to estimate the harm. Moreover, the model is not limited to situations, like Example 1, where damage is constant. There are important examples of indivisible harms that can be produced by concurrent causes, such as injuries to individuals from malpractice involving medical teams or personal injuries arising from multivehicle accidents. However, in many cases, the actual harm and the probability of harm may increase with the number of tortfeasors. It remains possible to show that under certain conditions, a tortfest can still occur when the expected harm increases with the number of actors (i.e., there is an increase in the benefit from the activity that is sufficiently greater than the increase in (diluted) liability from the increased harm). Finally, it is not claimed that tortfests are always welfare enhancing (they are not). See Dillbary, supra note 21, at 995.

29 See supra notes 10–11 and accompanying text.
any case because of the other factories. Example 1 shows why this argument is faulty. One or two factories acting together would engage in the harmless activity, as liability makes the tortious activity less profitable. However, the same two factories will switch to the tortious alternative activity if joined by a third factory. Each factory’s tortious behavior is, therefore, a but-for cause of the harm because no factory would have engaged in the tortious activity of producing cement without the third, and the victim would not be injured as a result.

Example 1 also illustrates that a non-injuring party can be a but-for cause of the harm. For example, assume that the alternative activity is drag racing involving two drivers and a spectator—the classic example in concerted-action cases. As with the factories, neither one of the drivers would participate in the race unless joined by the two other actors. This is because a group causation theory like concerted action would hold liable the injuring driver and the non-injuring parties (the spectator and the driver who did not hit the victim). As with the factories example, it is only once there are three participants in the tortious activity that it becomes the profitable choice for any of them. As a result, each of the non-injuring actors (the non-injuring driver and the spectator) is a but-for cause of the harm because but for their decision to engage in the tortious activity, the actor who physically

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30 See Dobbs, supra note 7, § 171 (discussing the two-fire examples and noting that substantial factor has also been extended to cases where each of multiple polluters contributed to the victim’s harm); Keeton et al., supra note 5, § 41 (arguing that the but-for test “fails” when multiple sufficient forces concur to bring about the harm); see also supra note 5.

31 The conduct of each actor can be considered a but-for cause of the harm even when four or more actors engage in the activity. For an in-depth discussion, see generally J. Shahar Dillbary, Causation Actually, 51 GA. L. REV. 1, 34–39 (2016), which shows that in many cases involving multiple actors, each can be a but-for reason for the victim’s harm, and providing policy justifications for applying a but-for presumption in other cases.


33 See, e.g., Wilson v. Firestone Tire & Rubber Co., No. 71981, 1985 Mich. App. LEXIS 3192, at *5 (Jan. 4, 1985) (“Even if defendant caused no harm himself, he is liable for the harm caused by his fellows because all acted jointly.”); Restatement (Second) of Torts § 876 cmt. d (Am. L. Inst. 1979) (noting that an actor who provided encouragement or assistance is considered a tortfeasor and is responsible for the resulting harm).
injured the victim would not engage in the tortious activity either, and the victim would not be harmed. While there is theoretical potential for group liability to incentivize torts because of its dilutive effect, our example makes it clear that this depends on what actors believe and choose in this liability environment. If they expect others to be deterred, then choosing the harmless activity is best. If they believe that the latent cost-dilution incentive will lead enough actors to commit the tort, then that becomes the best choice. There are thus two possible equilibria: deterrence or encouragement of the harmful activity. To date, a deterrent effect of group liability has simply been assumed.

II. THE EXPERIMENT

This Article’s goal is to empirically test how subjects behave in situations like Example 1. More specifically, our goal is to test whether: (1) imposing liability with the aid of group causation theories encourages actors to unconditionally engage in tortious activities instead of deterring them; (2) imposing liability may result in a tortfest dynamic (i.e., more actors would be willing to commit a tort as the number of actors increases); (3) the but-for test applies in settings like concurrent causes where courts insist it does not; (4) the nature of the activity and perceptions of its social value (e.g., producing a social good versus harming someone to get only a personal benefit) may impact willingness to engage in tortious activity; and (5) perceptions of fairness may change with the number of actors (e.g., a tortious activity may be considered more equitable if conducted by many). With the theory and goal in mind, we can now turn to the next step—the study’s design and its relation to the prior literature.

A. The Prior Literature

We begin with a review of the empirical literature. Previous attempts to measure the deterrent effects of tort law doctrines can be roughly divided

34 Indeed, but for the spectator’s decision to observe the race with approval, neither driver would agree to engage in the race, and no one would be harmed. Similarly, but for the non-injuring driver, the other driver would not drive carelessly even if encouraged by the spectator. For a discussion of the model’s assumptions and limitations, see Dillbary, supra note 31, at 57–63.

35 See, e.g., Posner, supra note 8, at 212 (discussing a scenario similar to Summers and noting that “it would be an economic mistake to let both off scot-free, as that would reduce deterrence”); see also supra notes 3–4, 6, 8–9, 17–18, 20.

36 See, e.g., Boim v. Holy Land Found. for Relief & Dev., 549 F.3d 685, 697 (7th Cir. 2008) (discussing concurring-causation cases and noting that in such cases, “the requirement of proving causation is relaxed”); Shackil v. Lederle Lab’yrs, 561 A.2d 511, 515 (N.J. 1989) (explaining that concert of action and alternative liability are exceptions to the “indispensable” cause-in-fact requirement).
into two types: (1) observational approaches and (2) experimental studies. To the best of our knowledge, none of these studies have directly tested the effect of group causation theories and the effects of dilution of liability nor have they tried to determine how an actor’s decision about whether or not to engage in a tortious activity impacts others who face the same dilemma.37

1. The Observational Approach

A number of prior empirical studies have used observational data to assess the deterrent effects of tort law, for example by comparing the outcomes from variations in liability regimes across jurisdictions and different types of tortious activities.38 In general, these studies provide mixed

37 There is a large, more general literature in psychology that investigates the relationships between causality, counterfactual reasoning, and attribution of responsibility. However, this literature does not directly engage the law’s structure and impacts as we do. See, e.g., Mark D. Alicke, Culpable Control and the Psychology of Blame, 126 PSYCH. BULL. 556, 571 (2000); David R. Mandel, Counterfactual and Causal Explanation: From Early Theoretical Views to New Frontiers, in THE PSYCHOLOGY OF COUNTERFACTUAL THINKING 11 (David R. Mandel, Denis J. Hilton & Patrizia Catellani eds., 2005) (reviewing psychological literature on causal explanations and attribution and exploring the role of counterfactual thinking); Denis J. Hilton, John L. McClure & Ben R. Slugoski, The Course of Events: Counterfactuals, Causal Sequences and Explanations, in THE PSYCHOLOGY OF COUNTERFACTUAL THINKING, supra, at 44 (exploring relationship between counterfactuals, conditions, and construction of specific and general causal explanations); Barbara Spellman, Alexandra P. Kincannon & Stephen J. Stose, The Relation Between Counterfactual and Causal Reasoning, in THE PSYCHOLOGY OF COUNTERFACTUAL THINKING, supra, at 28 (describing the relationship between counterfactual and causal reasoning generally, including the role of underlying knowledge and information, the contingent relations between different types of judgment, and the potential for causal attribution in certain cases without need for outcome-changing counterfactuals (but-for relationship)). See, e.g., Jill R. Horwitz & Joseph Mead, Letting Good Deeds Go Unpunished: Volunteer Immunity Laws and Tort Deterrence, 6 J. EMPIRICAL LEGAL STUD. 585, 586 (2009) (examining tort liability for volunteers across states and finding a deterrent effect); Frank A. Sloan, Emily M. Stout, Kathryn Whetten-Goldstein & Lan Liang, Drinkers, Drivers, and Bartenders: Balancing Private Choices and Public Accountability, 11 J. OF ECON. & POL’Y 28 (2017) (examining the relationship between alcohol availability and drinking behaviors across states and finding that tort liability has a deterrent effect). A number of studies investigated the deterrent effect of tort liability compared to no-fault auto-accident systems. Compare, e.g., Alma Cohen & Rajeev Dehejia, The Effect of Automobile Insurance and Accident Liability Laws on Traffic Fatalities, 47 J. OF EPIDEMIOLOGY & EPIDEMIOLOGICAL STUDIES 357, 382 (2004) (finding that introduction of no-fault rules led to an increase in fatalities), with Paul Zador & Adrian Lund, Re-analyses of the Effects of No-Fault Auto Insurance on Fatal Crashes, 53 J. RISK & INS. 226, 235 (1986) (finding the opposite effect). For further analysis of the impact of various tort reforms on auto accidents, see Paul H. Rubin & Joanna M. Shepherd, Tort Reform and Accidental Deaths, 50 J. OF L. & ECON. 221, 235 (2007), which analyzes the impact of various tort reforms, including damage caps, for auto accidents, and W. Jonathan Cardi, Randall D. Penfield & Albert H. Yoon, Does Tort Law Deter Individuals? A Behavioral Science Study, 9 J. EMPIRICAL LEGAL STUD. 567, 573–74 (2012), which provides a concise overview of these and other auto-accident studies. Scholars have also tried to empirically test the impact of tort liability on medical malpractice. E.g., Bernard S. Black, Amy R. Wagner & Zenon Zabinski, The Association Between Patient Safety Indicators and Medical Malpractice Risk: Evidence from Florida and Texas, 3 AM. J. HEALTH ECON. 109, 135 (2017) (finding that increased patient safety was correlated to fewer malpractice claims, such that the risk of suit could incentivize increased quality of care); Joanna M. Shepherd, Tort Reforms ‘
evidence on the deterrent effects of tort law. As Professors Theodore Eisenberg and Christoph Engel explain, such studies often suffer from a number of limitations. They are unable to “fine tune dependent and independent variables”; are often subject matter-specific (e.g., focusing on medical malpractice or car accidents); generally cannot observe decisions at the individual level, the dynamic of group activities, or how the actions of one group member impact the behaviors of others; and may suffer from reverse causality. Most importantly, observational studies are often not well suited to determine causal links. As a recent study explains: “[E]xisting deterrence studies all suffer from the same limitation—they use proxies as a means of measuring the causal link between tort liability and changes in tortious behavior.” Using proxies exacerbates another problem that is common to observational studies: the difficulty of determining whether the effects observed are the result of the legal rule investigated, some unobserved factors, or the unique characteristics of the dataset used.

Our innovative experimental design allows us to overcome these shortcomings and directly test the impact of causation law on actors’ behaviors. By employing a well-known technique called the Strategy

Winners and Losers: The Competing Effects of Care and Activity Levels, 55 UCLA L. REV. 905, 970 (2008) (finding that some medical malpractice tort reforms led to reduced tort awards and increased death rates, particularly for women); Michelle M. Mello & Troyen A. Brennan, Deterrence of Medical Errors: Theory and Evidence for Malpractice Reform, 80 TEX. L. REV. 1595, 1598 (2002) (surveying literature that studied the deterrent effect of tort and malpractice litigation).

39 Theodore Eisenberg & Christoph Engel, Assuring Civil Damages Adequately Deter: A Public Good Experiment, 11 J. EMPIRICAL LEGAL STUD. 301, 302 (2014) [hereinafter Eisenberg & Engel, Public Good Experiment] (discussing some limitations of studies).

40 Id. (stating that because of difficulty “specifying models of outcomes, such as fatality rates or damages awards,” researchers are “at the mercy of” observed variation).


42 See, e.g., Cardi et al., supra note 38, at 574 (noting limitations of tort-deterrence literature including a narrow focus on particular rules and contexts and the correlational nature of studies); Eisenberg & Engel, Public Good Experiment, supra note 39, at 302 (explaining that a lack of panel data generally prevents dynamic study of an individual’s response to tort liability and highlighting issues of omitted variables and reverse causality).

43 Cardi et al., supra note 38, at 570–71 (emphasis added) (using a survey methodology to analyze the effect of different legal regimes on actors’ willingness to commit a tort and concluding that unlike the threat of criminal liability, the effect of potential tort liability did not have a significant effect on subjects’ stated willingness to engage in risky behaviors).
Method, we analyze actors’ decisions at the individual level. Importantly, the Strategy Method allows us to observe whether a subject’s decision to behave tortiously is conditional upon others acting tortiously as well. It is thus especially well suited to analyze the law’s causal effects and the tortfest dynamic.

2. The Experimental Approach

Our design also differs from previous experimental approaches, which are themselves a recent phenomenon in torts scholarship. Some experimental studies use a survey-based approach. An example is the recent study by Professors W. Jonathan Cardi, Randall Penfield, and Albert Yoon, which asked first-year law students to rate the likelihood that they would engage in certain tortious activities and found no deterrent effect for tort law. Each student completed a survey that experimentally varied one of four possible legal regimes: no liability, criminal liability, tort liability, and unspecified liability. Survey studies are useful, but as the Cardi et al. study explicitly admits, they suffer from a major drawback that makes them unsuitable to test our hypotheses: survey studies rely on subjects’ self-reported predictions rather than their actual behavior. The concern is that when it comes to making the decision in question—e.g., whether or not to commit a tort—subjects would behave differently than they said they would. Moreover, survey studies are often static in nature and do not allow

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44 See infra Section II.B.4 (explaining the Strategy Method).
45 See, e.g., Theodore Eisenberg & Christoph Engel, Unpacking Negligence Liability: Experimentally Testing the Governance Effect, 13 J. EMPIRICAL LEGAL STUD. 116, 120 (2016) [hereinafter Eisenberg & Engel, Unpacking Negligence Liability] (explaining that while “[i]n economics[,] the experimental method has become very popular because experiments directly test formal theoretical models,” surprisingly, the “formal economic models of legal institutions have not very often been tested in the lab,” and noting that “[t]here is no more than a handful of pertinent experimental papers”).
46 See Cardi et al., supra note 38.
47 Id. at 578. In the unspecified regime, no legal regime was given, and “subject[s] [were] left with [their] own understanding of the possibility of sanctions or liability.” Id. at 578–80.
48 Id. at 571 (explaining that their experiment “engenders its own limitations,” including the fact that it “captured only participants’ self-predicted behaviors, not their actual behaviors”). The potential gap between self-reported choices and actual choices is a limitation common to experimental survey designs.
49 For an example of studies that compare laboratory experiments and survey studies, see Edward L. Glaeser, David I. Laibson, José A. Scheinkman & Christine L. Soutter, Measuring Trust, 115 Q.J. ECON. 811, 841 (2000), which found standard attitudinal survey questions about trust are not very accurate at predicting trusting behavior in their experiments (in terms of trusting others to reciprocate monetary transfers in an incentivized trust game). Also, see Christoph Bühren & Thorben C. Kundt, Imagine Being a Nice Guy: A Note on Hypothetical vs. Incentivized Social Preferences, 10 JUDGMENT & DECISION MAKING 185, 185 (2015), which determined that behavior in dictator games is different between incentivized and hypothetical games.
researchers to observe the decision dynamic in real time, time trends, or, importantly, whether one’s action is dependent on her belief as to whether and how many actors will behave tortiously.  

Like the Cardi et al. study, our methodology tests the causal impact of tort liability by comparing results for regimes with liability and no liability, as well as across contextualized scenarios and neutrally framed ones. However, our design is different in several important aspects. Chief among them is that we observe the subjects’ actual behavior in a context with real financial consequences, rather than relying on self-reports. Moreover, using a between-subject framework with multiple decision rounds, we are able to observe subjects’ decisions in different liability environments over time. This dynamic environment allows us to investigate whether experience changes their decisions or their beliefs about how others will respond.

Lab experiments are a more recent, dynamic form of experimental study in tort law. A number of researchers have used experiments to assess how variations in tort-liability regimes might affect behavior. In a well-known study, Professors Lewis Kornhauser and Andrew Schotter asked each of the 113 subjects to choose (at a cost) their level of care, which determined the probability that an accident would occur. Under a strict-liability regime, a subject paid a predetermined amount if the accident occurred. Under a negligence regime, payment was due only if the care taken was below an announced standard of care. The authors found the choice of liability regime had a significant effect. Professors Kornhauser and Schotter’s experiment “consider[ed] the simplest possible accident model” in which a “single” actor chose her desired precaution level. Their experiment did not incorporate a victim, consider the tortfeasor’s choice of the level of activity, or involve interactions between multiple potential tortfeasors. Subsequent work by Professors Vera Angelova, Giuseppe Attanasi, and Yolande Hiriart.

50 See, e.g., Cardi et al., supra note 38, at 576–77 (administering a survey as a single event exposing participants to randomized alternate vignettes to elicit causal relationship); Eisenberg & Engel, Public Good Experiment, supra note 39, at 302 (noting advantages of experiments: “subjects repeatedly interact over time,” making it possible to observe “the dynamic individual-level process of behavior” and “observations over time enable evaluating the process leading to observed outcomes”).

51 Lewis Kornhauser & Andrew Schotter, An Experimental Study of Single-Actor Accidents, 19 J. LEGAL STUD. 203, 204, 208–09 (1990) (analyzing liability effects in a “model in which the actions of a single individual determine the expected frequency of occurrence of a loss”). The authors found that under a negligence regime, individuals chose the (exogenous) reasonable level of care. By contrast, under the strict-liability regime, the investment in care initially exceeded, then fell below, the efficient level. Id. at 231–32.

52 Id. at 208–09.

53 Id. at 213 (noting the hypothesis tests’ significance at 5% level).

54 Id. at 204.
incorporated a victim as part of the experiment, as well as an additional baseline treatment with no liability.55

Building on this work, Professors Eisenberg and Engel designed an investment task where the more the active party invests, the higher the risk of harm is to another (a passive innocent party).56 They found that tort law has an expressive function that strengthens its deterrent effect by creating nonpecuniary motivations to avoid harming others. 57 This expressive function may result from the distaste for harming an innocent party, the tendency to adhere to social norms, and “the desire to avoid blame.”58 We follow Eisenberg and Engel’s steps by recognizing the possibility that psychological (nonpecuniary) costs may slow down the occurrence of a tortfest.59 However, unlike Eisenberg and Engel, our design also allows us to investigate how the dynamic interaction between prospective tortfeasors and the impact of shared liability change fairness perceptions. We find some support for the proposition that psychological costs (to the extent they are correlated with fairness) may be higher in the case of a single tortfeasor, but that they monotonically decrease as the number of tortfeasors increases—a phenomenon Eisenberg and Engel could not observe due to their design.

A few researchers have used public-goods-style experiments to assess the impacts of tort law.60 In a different study, Eisenberg and Engel used a variant of a public-goods experiment to study the impact of tort-damage regimes.61 However, their study did not consider the potential for shared liability and joint action by wrongdoers. The more recent study by Professors Bruno Deffains, Roman Espinosa, and Claude Fluet used a public-goods

56 Eisenberg & Engel, Unpacking Negligence Liability, supra note 45, at 119.
57 Id. at 134 (noting deterrent effects from courts that “make[] social expectations explicit” and “blame” defendants who fail to meet them, as distinct from incentives created by pecuniary liability).
58 Id. at 116.
59 Id. at 139 (finding that nonpecuniary expressive dimensions enhance deterrence). But note that in Eisenberg and Engel’s design, pecuniary liability is not contingent on the actions of other players, so there is no tortfest dynamic. We extend consideration of their findings by considering the implications in our design.
60 Public-goods experiments generally involve participants making private decisions about contributions to produce a public good that generates benefits which can be shared by all, even those who do not contribute fully. See, e.g., Rachel T.A. Croson, Public Goods Experiments, in BEHAVIOURAL AND EXPERIMENTAL ECONOMICS 221, 221–22 (Steven N. Durlauf & Lawrence E. Blume eds., 2010).
61 Eisenberg & Engel, Public Good Experiment, supra note 39, at 331 (contrasting individual tort damages with class action and punitive-damage analogs and finding some impact in the latter two categories).
experiment to investigate the impact of liability rules in cases where participants face trade-offs between privately profitable actions and imposing harm on others.\(^{62}\) Similar to Eisenberg and Engel, the Deffains et al. study found that liability rules have a “norm-activation effect.”\(^{63}\) They convey a message that leads participants to put more weight on social concerns and harms to others thereby increasing deterrence.\(^{64}\) The effect appears strongest for strict liability, and it complements the impact of social norms, which the authors believe are operative even when legal sanctions are absent or imperfectly enforced.\(^{65}\) Like its predecessors, however, the study did not consider the impact of liability in the context of multiple tortfeasors or the potential for shared liability.

### B. The Experimental Design and Procedure

To our knowledge, this Article is the first to use an experimental approach to test the impact of shared tort liability on actors’ decisions whether or not to engage in a tortious activity while observing the subjects’ individual preferences. To directly test the effect of group causation theories, we designed a series of experiments, based on Example 1, in which actors independently and simultaneously decide whether or not to engage in a tortious activity, knowing others face the same dilemma. Importantly, in our experiments, acting is worthwhile only if a minimum number of actors decide to engage in the tortious activities. As such, our experiments can be viewed as a form of a “stag hunt” game.\(^{66}\) Broadly defined, these are cases

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\(^{63}\) Id. at 2.

\(^{64}\) Id. at 2, 14.

\(^{65}\) Id. at 2–3, 14. Note that the finding of a stronger effect under strict liability versus negligence was supported only by “weak evidence.” Id. at 14.

\(^{66}\) The logic of a “stag hunt” game is as follows: Let us suppose that there are a number of hunters who must choose between putting effort toward hunting a hare or a stag (i.e., a deer). The chances of catching a hare are independent of what others do. There is no chance of catching a stag alone, but the chances of a successful stag hunt go up sharply with the number of hunters who decide to join the hunt. A stag is much more valuable than a hare. Participants must decide independently whether to gamble the sure gain of the hare for the chance of success at a stag. In our experiment, the tort is similar to a “stag” in that committing the tort can be more beneficial than simply keeping the endowment, but it only becomes so with more tortfeasors acting. For a general discussion of such coordination games, see, for example, David Schmidt, Robert Shupp, James M. Walker & Elinor Ostrom, Playing Safe in Coordination Games: The Roles of Risk Dominance, Payoff Dominance, and History of Play, 42 GAMES & ECON. BEHAV. 281, 281 (2003). Our design is also related to threshold public-goods experiments, where a certain level of cooperation is needed for the production of the public good. In most cases,
in which parties’ payoffs depend on the level of implicit cooperation among subjects. Parties do not communicate in any manner with others about their strategies and choices. Rather, cooperation between potential tortfeasors can be achieved through trust or based on belief (in the case of unconditional decisions\textsuperscript{67}) that others will behave in a certain way.

Our experiments provide a consistent measure of the hypotheses tested and allow the study of repeated individual behavior over time. Importantly, they allow us to observe individual-level decision-making and, using the Strategy Method, to see how actors react causally to different incentive schemes.\textsuperscript{68} In the experiments, we manipulate variables of interest (e.g., whether there is tort liability or not) while subjects make choices that expose them to different actual monetary payoffs, dependent on the actions of other subjects, with the potential to impose real losses on a party in the experiment. For these reasons, this approach is especially well suited to test group dynamics in cases when the actions of one tortfeasor can benefit others through doctrines of causation that result in shared liability.\textsuperscript{69}

To test our hypotheses, we presented subjects with four incentive schemes (treatments) similar to the one described in Example 1 (reproduced below for convenience):

**Example 1.** A number of actors\textsuperscript{70} (e.g., Factories) must choose between two activities: (1) a harmless activity (e.g., producing limestone) or (2) an alternative tortious activity (e.g., producing cement—a necessary ingredient for construction—that causes pollution). The harmless activity promises a $24 benefit whereas the alternative activity promises $35. However, if one or more actors engage in the alternative activity, an innocent victim will incur an expected harm of $24. The alternative activity is subject to liability. Taking care is either impossible or too costly.

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\textsuperscript{67} The use of unconditional and conditional decisions using the Strategy Method is discussed below. See infra Section II.B.4.

\textsuperscript{68} For an explanation of the Strategy Method, see infra Section II.B.4.

\textsuperscript{69} See infra Section II.B.4. For the external validity of lab experiments, see infra note 158 and accompanying text.

\textsuperscript{70} We use the term “actors” to refer to potential tortfeasors—that is, to those who have a choice and must decide whether to engage in a tortious activity. In the Context-Free frame, these actors are referred to as “Type-X” subjects. In the Vignette frame, they are referred to as the “Factories.”
1. The Context-Free (Neutrally Framed) Treatments

In the Context-Free treatments, the experiment was framed in a way that is neutral in the sense that it does not include references to a “real world” context that might influence participants’ perceptions. This neutral framing is common in experimental studies as a way to focus on the basic underlying payoffs and how participants respond.\(^71\) We consider it first, before moving to an experiment that describes the context in a way that better reflects the choices facing potential tortfeasors in our Vignette treatments. Subjects participated in groups of five in a repeated setting (i.e., rounds). In each round, each subject was endowed with 24 experimental tokens. Four of the five subjects, called Type-X, were assigned as actors,\(^72\) and one subject, called Type-Y, was assigned as a party who may incur a loss. Type-X subjects, analogues of the Factories in Example 1, decided whether or not to “act” (i.e., whether or not to engage in the harmful activity). The subjects who chose to act lost their 24-token endowment but in return received 35 tokens—the higher payoff from the harmful activity. The one Type-Y subject in the group, an analogue of the potential victim in Example 1, did not make any decisions during the experiment.\(^73\) However, Type-Y’s payoffs depended on the decisions of the Type-X subjects. If at least one Type-X subject chose to act, the Type-Y subject would have lost her 24-token endowment (i.e., the injury from the activity).

The neutrally framed treatments varied in the applicable liability regime: Liability versus No Liability for the actors. In the Context-Free–Liability treatment (CF–L), all Type-X subjects who chose to act had to fully compensate the Type-Y subject in equal shares.\(^74\) For instance, if two Type-X subjects chose to act, each lost her 24-token endowment—the opportunity cost of forgoing the harmless activity—and received the payoff from the harmful activity of 35 tokens. Each was then required to transfer 12 tokens \((24/2)\) to the Type-Y subject for a total gain of 23 tokens \((35 – (24/2))\). If three Type-X subjects chose to act, each gained 27 tokens \((35 – (24/3))\).


\(^{72}\) See supra note 70.

\(^{73}\) All groups, across all treatments, included a victim who was an actual participant. The victim is referred to as “Type-Y” in the Context-Free (neutrally framed) treatments and the “Owner” in the Vignette treatments. We allocated this role to a participant for two main reasons. The first was to avoid deception. Because subjects were told that their actions would result in a harm to a third party, the potential victim had to be included. Second, the knowledge that the victim is real and present can impact subjects’ decisions.

\(^{74}\) This apportionment rule is similar to that of several liability or joint and several liability with contribution. Because all subjects received an initial endowment, each is solvent.
In the Context-Free–No Liability treatment (CF–NL), none of the Type-X subjects who chose to act were required to compensate the Type-Y subject. Thus, if at least one of the Type-X subjects chose to act, the Type-Y subject simply lost her 24-token endowment. For example, if two Type-X subjects chose to act, each received 35 tokens from acting, and neither was required to transfer tokens to the Type-Y subject.

This means that under the No Liability treatment, the monetary payoffs for Type-X subjects who chose to act were constant and equal to 35 tokens. By contrast, under the Liability treatment, the monetary payoffs of Type-X subjects who chose to act were dependent on the total number of Type-X subjects who chose to act. As Table 1A shows, in the Liability treatment a Type-X subject was better off acting only if she was joined by two or three additional Type-X subjects (27, 29 > 24) but otherwise was better off keeping her 24-token endowment and refraining from acting (11, 23 < 24).75

Table 1A: The Incentives to Act as a Function of the Number of Actors

<table>
<thead>
<tr>
<th>The Number of Acting Type-X Subjects</th>
<th>Payoffs of Each Acting Type-X Subject</th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>24 (= endowment)</td>
</tr>
<tr>
<td>1</td>
<td>11 (= 35 – 24/1)</td>
</tr>
<tr>
<td>2</td>
<td>23 (= 35 – 24/2)</td>
</tr>
<tr>
<td>3</td>
<td>27 (= 35 – 24/3)</td>
</tr>
<tr>
<td>4</td>
<td>29 (= 35 – 24/4)</td>
</tr>
</tbody>
</table>

Note that although all Type-X subjects faced the same monetary incentive scheme, their actual gains from acting likely varied. The reason is that subjects may hold heterogeneous nonmonetary valuations associated with their actions. Notions of fairness, morality, inequality aversion, social preferences, or other-regarding preferences, to name a few determinants, may impact the actors’ subjective payoffs and, accordingly, their decisions.76

75 When a subject is joined by two actors—that is, as part of a group of three tortfeasors—each gain 35 tokens and must pay one-third of the damage, or 8 tokens (24/3), for a net gain of 27 tokens. When a subject is joined by three actors, each pays less—only one-fourth of the damage for a net gain of 29 tokens (35 – (24/4)). In both cases, acting is worthwhile as it promises a higher gain compared with not acting (27, 29 > 24). By contrast, acting alone or with one other actor promises less than the 24-token endowment. Acting alone gives the actor 35 tokens, but after paying the victim 24 tokens for her damages, the actor has a net gain of only 11 tokens (35 – 24). Acting with another means that the two actors will share the cost of the victim injury. In such a case each will gain 23 tokens (35 – (24/2)), which is still less than the initial 24-token endowment.

For example, some subjects could be comfortable knowing that if they act, the Type-Y subject would lose her endowment. Others might experience some nonmonetary, psychological cost from gaining at the expense of another, regardless of whether the Type-Y subject is compensated. In general, we would expect these nonmonetary costs to be relatively small in our experimental setting, as participants interact anonymously.

2. The Vignettes

In addition to the Context-Free treatments, we also conduct Vignette-based treatments. The Vignettes provide a chance to check the internal validity of our results in a contextualized, real-world setting with the same incentive structure as our Context-Free treatments. They also help bridge the gap between our experimental setting and the real world, supporting the practical relevance of our results.

In order to preserve comparability with our Context-Free treatments, we adopted a Vignette involving harm to property, potentially caused by a “polluting” factory. The liability regimes and payoffs remained the same as in the Context-Free treatments; the only difference was in the framing of the instructions. These additional treatments used Vignettes to frame the decisions in a more realistic setting. In the Vignette–Liability (V–L) and Vignette–No Liability (V–NL) treatments, four of the subjects—the actors who had to decide whether to engage in the tortious activity—were referred

77 For discussion of an example and the general the impact of nonmonetary costs on actors’ willingness to engage in harmful activity, see infra notes 95–99 and accompanying text.

78 See Elizabeth Hoffman, Kevin McCabe, Keith Shachat & Vernon Smith, Preferences, Property Rights, and Anonymity in Bargaining Games, 7 GAMES & ECON. BEHAV. 346, 347–48 (1994) (finding that anonymity in bargaining games increased subjects’ self-regarding decisions, suggesting other-regarding behavior is due to a desire for reciprocity instead of a preference for fairness); Gary Charness & Uri Gneezy, What’s in a Name? Anonymity and Social Distance in Dictator and Ultimatum Games, 68 J. ECON. BEHAV. & ORG. 29, 34 (2008) (finding that increasing social distance from other subjects encouraged self-interested behavior in bargaining games).

79 The property context allowed us to control the financial incentives faced by subjects and keep any psychological costs comparable to our Context-Free variant, in which “victims” lost only money. If, for example, we had used a personal-injury vignette, some subjects could have (hypothetically) expected some nonmonetary harms and may also have had very different psychological or moral costs associated with such harm. Although the context of “polluting” factories might be considered politically sensitive, we chose neutral vignette language—for example, we do not use the terms “pollute” or “victim” in the Vignettes. The effect we find is consistent across both our neutral and contextualized framings, which further suggests it is not simply an artifact of the choice of Vignette frame.

80 The difference in framing between the Context-Free treatment and the Vignette is in the terms used to describe the actors and the harmful activity. Type-X subjects, Type-Y subject, and “acting” in the former were replaced with “Factories,” “Owner,” and “switching to cement production” in the latter. The difference in framing is summarized in Table 1B.
to as Factories. The fifth subject—the potential victim—was referred to as the Owner of another structure nearby. Each Factory chose whether to keep producing limestone—“not acting” in the Context-Free treatments—or switch to producing cement—“acting” in the Context-Free treatments. If a Factory produced limestone, it kept its 24-token endowment. If a Factory switched to producing cement, the tortious activity, it lost its endowment but in return earned 35 tokens from producing cement. However, if at least one Factory switched to producing cement, the Owner nearby lost 24 tokens—the full value of her structure.

Table 1B summarizes the differences in framing across treatments. These changes to the frame of the instructions did not alter the monetary incentives of the decision environment. They may, however, have changed subjects’ perceptions. The contextual frames of the Vignettes allow us to test the robustness of our results from the neutrally framed experiments in this more realistic setting.

<table>
<thead>
<tr>
<th>Table 1B: The Difference in Framing Across Treatments</th>
</tr>
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<tbody>
<tr>
<td><strong>Context-Free</strong></td>
</tr>
<tr>
<td>(CF–L, CF–NL)</td>
</tr>
<tr>
<td><strong>Actors</strong></td>
</tr>
<tr>
<td><strong>Potential Victim</strong></td>
</tr>
<tr>
<td><strong>Decision</strong></td>
</tr>
</tbody>
</table>

For brevity and consistency when discussing the treatments, we refer collectively to choices to act—in the Context-Free treatments—or produce cement—in the Vignettes—simply as the “choice to act.” Declining to act or producing limestone are both referred to as the “choice not to act.”

3. **Sessions and Rounds**

Each of the four treatments (summarized in Table 3 in Section II.E below) was administered in multiple separate sessions during which the treatment was repeated in seven rounds. At the beginning of each session, subjects were randomly assigned to one of two types: Type-X or Type-Y in the Context-Free treatments and Factory and Owner in the Vignettes. In each treatment, subjects in both roles—Type-X and Type-Y—received the same

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81 Contextualized instructions in economics experiments can, in some instances, change behavior compared to behavior with neutral instructions. See Alekseev et al., supra note 71, at 49 (“[W]hen instructions are embedded in a meaningful, as opposed to abstract, context, behavior may change.”); Leigh Raymond & Timothy N. Cason, Can Affirmative Motivations Improve Compliance in Emissions Trading Programs?, 39 POL’Y STUD. J. 659, 673 (2011) (finding that a neutral frame inspired less compliant behavior compared to context-framed treatments).
Subjects assigned as Type-X or Factories made their decisions in each round simultaneously—that is, without any knowledge of the other actors’ decisions in the given round. At the beginning of each round, groups were randomly rematched (i.e., Strangers Matching) to mitigate any reputational effect from repeated play with the same individuals. However, a subject’s type—Type-X or Type-Y—remained constant throughout the entire session. This avoided any potential effects that might have arisen from the order in which subjects played the two roles. At the end of each round, all subjects received information on the number of actors who chose to act and the final token earnings for the round. Additionally, actors were notified about their payoffs, and the potential victim was notified whether she incurred any losses and whether she was compensated. All subjects’ payoffs were a function of the liability regime and the number of subjects who chose to act.

4. **The Strategy Method**

To test the applicability of the but-for test in settings like concurrent causes and whether and under what conditions a tortfest would develop, we employed the Strategy Method. This method is used frequently to analyze subjects’ “willingness to contribute to public goods conditional on others’ contributions,” and is therefore especially suited for our purposes. This

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82 To avoid influencing subjects, the instructions did not use terms such as actor, aggrieved party, tortfeasor, victim, or tort. See Alekseev et al., supra note 71, at 57 (explaining that the use of “evocative” terms may elicit emotional responses from subjects).

83 In an experimental session, subjects can be matched into fixed groups across decision rounds (i.e., Partners Matching) or subjects in the session can be randomly rematched into new groups across decision rounds (i.e., Strangers Matching). See generally James Andreoni & Rachel Croson, Partners Versus Strangers: Random Rematching in Public Goods Experiments, in 1 HANDBOOK OF EXPERIMENTAL ECONOMICS RESULTS 776 (Charles R. Plott & Vernon L. Smith eds., 2008) (discussing relevant studies on the effect of Partners Matching versus Strangers Matching).

84 For examples of studies that have utilized the Strategy Method, see Urs Fischbacher, Simon Gächter & Ernst Fehr, Are People Conditionally Cooperative? Evidence from a Public Goods Experiment, 71 ECON. LETTERS 397, 398 (2001), which employed a variant of the Strategy Method to examine conditional cooperation, and Urs Fischbacher & Simon Gächter, Social Preferences, Beliefs, and the Dynamics of Free Riding in Public Goods Experiments, 100 AM. ECON. REV. 541, 543–44 (2010), which applied the Strategy Method to test subjects’ contribution preferences. For a recent survey of the literature regarding the effectiveness of the method, see Jordi Brandts & Gary Charness, The Strategy Versus the Direct-Response Method: A First Survey of Experimental Comparisons, 14 EXPERIMENTAL ECON. 375, 394–95 (2011), which reported that, in a few instances, employing the Strategy Method compared to the “direct-response” method (i.e., only making unconditional decisions) may have impacted subjects’ behaviors, but arguing that their study “should at least dispel the impression that the strategy method inevitably yields results that differ significantly from results gathered using the traditional direct-response method.”

method allows us to observe whether a subject’s decision to act is conditional on others’ decisions to do the same. For example, it allows us to observe whether a Factory would prefer to engage in the tortious activity only if two or more other Factories would do the same but would prefer to engage in the harmless activity otherwise.

The method works as follows. In the beginning of each round, every actor—subjects assigned as Type-X or Factories—made two types of decisions. First, the actor made an unconditional decision to act. Following the unconditional decision, each actor made conditional decisions by filling out an Action Table in which they chose whether to act in all possible conditional scenarios. The Action Table in the Context-Free treatments are reproduced below as Tables 2A and 2B.

**Table 2A: Type-X Subject Unconditional Decision to Act**

<table>
<thead>
<tr>
<th>Your Unconditional Decision to Act</th>
<th>Act or Not</th>
</tr>
</thead>
<tbody>
<tr>
<td>Do you choose to act?</td>
<td></td>
</tr>
</tbody>
</table>

**Table 2B: Type-X Subject Conditional Decision to Act (Action Table)**

<table>
<thead>
<tr>
<th>Your Conditional Decisions to Act (Action Table)</th>
<th>Act or Not</th>
</tr>
</thead>
<tbody>
<tr>
<td>If none of the other three Type-X members choose to act, would you choose to act?</td>
<td>Act or Not</td>
</tr>
<tr>
<td>If one of the other Type-X members chooses to act, but the other two Type-X members choose not to act, would you choose to act?</td>
<td>Act or Not</td>
</tr>
<tr>
<td>If two of the other Type-X members choose to act, but the other Type-X member chooses not to act, would you choose to act?</td>
<td>Act or Not</td>
</tr>
<tr>
<td>If all three of the other Type-X members choose to act, would you choose to act?</td>
<td>Act or Not</td>
</tr>
</tbody>
</table>

At the end of each round, the Action Table of one of the four actors was randomly chosen for that round’s earnings calculation. For the other three

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86 See infra Table 2A.

87 See infra Table 2B.
actors in the group, only their unconditional decisions were used to calculate earnings in that round. The following example illustrates the procedure.

Example 2. Consider a Context-Free treatment with a Liability regime. Suppose Type-X subject X1 is randomly chosen in a round. This means the computer will use X1’s Action Table to calculate earnings in this round. The unconditional decisions for Type-X subjects X2, X3, and X4 will be used to calculate earnings. Suppose further that X2 and X3 made unconditional decisions to act and X4 made the unconditional decision not to act.

Suppose X1 indicated in her Action Table that she would act if two other Type-X subjects chose to act. In such a case, X2 and X3 (both of whom chose to act unconditionally) and X1 (who chose to act if two others joined her) would act in this round. As a result, X1’s, X2’s, and X3’s (monetary) earnings would each be 27 tokens \((35 - (24/3))\), whereas X4, who chose not to act unconditionally, would earn 24 tokens. Type-Y’s earnings (i.e., compensation) would be 24 tokens. The Strategy Method creates appropriate incentives for individuals to reveal their true preferences because of the chance that their choices, unconditional and conditional, could be used to determine their payoffs from the experiment.

5. Procedure: Review Questions, Risk Proclivity, and Exit Questionnaire

All sessions for the pilot and main treatments were conducted at Appalachian State University in the Appalachian Experimental Economics Lab (AppEEL) using z-Tree software. After entering the lab, subjects were randomly assigned to computer stations. After reading the instructions, subjects answered review questions to test their understanding of the basic protocol and incentives of the experiment. The first round began only after all review questions were answered correctly. Following the experiment’s seven rounds, subjects also completed the monetarily incentivized Holt and Laury risk-preference task to measure risk aversion.  

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88 In addition to their Types (Type-X or -Y), subjects were also assigned an anonymous ID each round. Type-X IDs were X1, X2, X3, and X4. The Type-Y ID was Y. In the contextually framed Vignette treatments, Factory IDs were F1, F2, F3, and F4, and the Owner ID was O.

89 The difference between the 35 tokens from acting and the 8 tokens \((24/3)\) that each will need to transfer to fully compensate Y for her 24-token loss.

90 See Charles A. Holt & Susan K. Laury, Risk Aversion and Incentive Effects, 92 AM. ECON. REV. 1644, 1645 (2002). This has become the standard method for assessing risk aversion in experimental settings; it has been cited over 6,000 times. See, e.g., Lisa R. Anderson & Jennifer M. Mellor, Are Risk
Subjects then completed a post-experiment questionnaire. In addition to demographic data on subjects’ characteristics, such as age and education, subjects were asked to rank the fairness of the decision to act—i.e., choose the harmful activity—as a function of the number of actors. Using a scale from 0 to 5, with 0 being unfair and 5 being completely fair, subjects evaluated the fairness to: (1) the victim, (2) the actors, and (3) society, defined as all of the subjects in the group (i.e., the actors and the victim).

To reduce the incentive to hedge across rounds, subjects were paid in cash at the end of each session. The payment was based on one randomly chosen decision round in the experiment. For example, if Round 4 was randomly chosen from the seven rounds, the subjects received a cash amount based on the number of tokens they earned in Round 4.

6. Subject Characteristics

Our subjects were drawn from the student population at Appalachian State University. Although our experimental method does not require a sample that is representative of the general population, we collected data on participant characteristics to check and control for whether any particular aspects of our subject pool influenced our results. A summary of our sample can be found in Table AII-6. Slightly more than half (56%) of the subjects were females. Approximately 78% of the subjects were white. The average age of our subjects was twenty-two. Slightly more than half of the subjects indicated belief in a higher power of some form (57%). Almost exactly half our subjects indicated a political orientation; among them approximately 58% were Democrats and 42% were Republicans. Across political parties, the average intensity in political party was 2.33 out of 5. Just under 43% of the subjects had majors in the Business School. On average, subjects had participated in one previous economics experiment in the lab. Our subjects were slightly risk averse, scoring 4.5 on average on the Holt and Laury scale, for which larger measures indicate more risk-seeking preferences, and a measure of 5 is risk neutral.

The collection of these participant controls allows us to see the variation we have in our sample of

Preferences Stable? Comparing an Experimental Measure with a Validated Survey-Based Measure, 39 J. RISK & UNCERTAINTY 137, 137–38 (2009) (recognizing the Holt and Laury lottery-choice task as the “‘gold standard’ in the experimental literature on risk aversion”). This measure allowed us to control for the possibility that differences in risk tolerance across participants could influence our results.

See infra Section II.E.5 for a discussion of regressions with controls, finding that none of the measured participant characteristics drive our results.

See infra Table AII-4 for a breakdown by liability regime. The scale is based on the number of instances a subject chooses Lottery B over Lottery A. Holt and Laury designed their lottery-choice task such that, in a menu of ten lottery choices—A and B—the expected value of Lottery B first becomes larger than Lottery A for Decision 5. If a subject chooses Lottery B more than five times, they are exhibiting risk-seeking preferences. See Holt & Laury, supra note 90, at 1645, 1649.
subjects and empirically test for the importance of these individual characteristics.

C. The Hypotheses and Theoretical Predictions

We used the various treatments and the Strategy Method to test the following hypotheses:

**Hypothesis 1: Reduced Deterrence.** The imposition of tort liability on multiple actors can encourage, rather than discourage, actors to engage in tortious activities. We hypothesize that the imposition of liability on multiple actors will incentivize an actor who would alone engage in a harmless activity to *switch* to a tortious activity if liability is imposed on a large enough number of actors. We test this hypothesis by comparing decisions to act in settings in the Liability regime and settings in the No Liability regime.

**Hypothesis 2: Tortfest.** The imposition of tort liability can result in a tortfest, where the more actors join the tortious activity, the more beneficial and enticing the activity becomes. We hypothesize that the demand for a tortfest will increase with the number of tortfeasors. We test this hypothesis by using the Strategy Method to compare individuals’ conditional decisions as the number of actors increases.

**Hypothesis 3: But-For.** The but-for test can be operable in tortious activities involving multiple actors, such as concurrent causes. In overdetermined cases, where each actor’s action can alone destroy the victim’s endowment, each concurring force is a but-for cause of the harm. This means that in the case of $n$ actors, but for the engagement of the marginal tortfeasor, none will engage in the activity, and the victim will not be harmed. This hypothesis is also tested by examining individuals’ conditional decisions using the Strategy Method.

**Hypothesis 4: The Nature of the Activity—Context.** We hypothesize that the context of a tortious behavior can affect the salience of costs and benefits, including any fairness concerns, making them more apparent to individuals. We test this hypothesis by contrasting the contextualized Vignette frame with the Context-Free frame. The factory Vignette uses a realistic scenario that highlights the social benefit from the tort (cement production). By contrast, in the Context-Free frame, the social benefits from the activity are less salient. Although the payoffs are the same in both treatments, we hypothesize
that the willingness to act will be higher in the Vignette frame compared to the Context-Free frame.

**Hypothesis 5: Alternative Fairness Ideals.** Tort liability is often linked to ideas of fairness. We have two competing hypotheses about the way tort liability may impact actors’ perceptions of fairness, depending on the number of actors:

**A. The Expressive Function and Moral Culpability.** One possibility is that, in addition to its consequent monetary costs, tort liability may signal the unfairness or moral culpability of an action. This would suggest individuals will be less willing to act under the Liability regime and perceive acting as less fair to the victim. The activity may be perceived as more morally reprehensible as the number of tortfeasors increases—that is, as more actors “gang up” on the victim.

**B. Moral Excuses and Mob Mentality.** Another possibility is that tort liability may mitigate perceptions of unfairness. For example, actors may see liability as a “price” for acting. This price may crowd out social norms and reduce any psychological costs from harming the victim, so they are more easily offset by monetary benefits. A larger number of acting participants may also imply that the tortious activity is socially acceptable or at least not as reprehensible as it would be perceived if only one or a few engaged in it. Because the monetary

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95 See, e.g., Bruno Defains & Claude Fluet, *Legal Liability when Individuals Have Moral Concerns*, 29 J.L. ECON. & ORG. 930, 942 (2013) (using theoretical treatment and finding crowding out of moral concerns with perfectly enforced tort liability). It is important to note that our particular experiments allow us to test this theory only in the context of harms to property, not tortious physical harm to individuals.

96 Here, it is important to note that in a tortfest dynamic, when the social benefit of the action changes from negative to positive, there may also be a corresponding effect on fairness perception: as more actors join in a “harmful” activity, the action may be perceived as more morally acceptable. For similar arguments related to the potential expressive role for law in shaping social norms, see supra note 93 and accompanying text.
price imposed by tort liability decreases with the number of tortfeasors, as liability is shared by all, a tortfest may also erode the psychological costs associated with the tortious activity when the activity is conducted by many.

We test these two competing hypotheses by comparing willingness to act across the liability treatments and by examining the fairness ratings in the post-experiment questionnaire.

We recognize that under the No Liability regimes, actors may incur a moral or psychological cost if they decide to act, as acting imposes an uncompensated loss on the innocent subject. Accordingly, a few observations regarding the No Liability treatments are in place. First, we expect that the psychological cost would be highest when an actor decides to act alone because that actor would be the sole reason for the harm. Consequently, the psychological cost may decrease as the number of actors increases, as moral responsibility is shared. A large number of tortfeasors may also imply that the activity is more socially acceptable and thus, by definition, less morally costly. Finally, the monetary benefits may simply outweigh any psychological costs. We expect that subjective psychological costs would not be especially high in our experimental setting, but subjects who have particularly strong moral intuitions may still decline to act because of these psychological costs.

In the Liability regimes, there is no dominant strategy in terms of expected monetary payoffs. With only one or two subjects acting, each is better off not acting, but with two or more subjects acting together, each is better off acting. There are therefore two possible pure-strategy Nash equilibria: no subjects choose to act or all four subjects choose to act. Subjects’ unconditional choices reflect their beliefs about the degree of implicit cooperation when the shared Liability rule is in place. The Strategy

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97 See, e.g., Deffains & Fluet, supra note 95, at 930–31 (discussing generally reluctance of individuals to impose harm on others due to costs in the form of guilt, in addition to possible social disapproval); id. at 938–39 (stating that in a No Liability regime, “individuals will take some precautions because . . . they anticipate the guilt of causing uncompensated harm”).

98 To illustrate, consider a No-Liability session in which each of the four Type-X subjects, X1–X4, must decide whether to act for a gain of 35 tokens or refrain from acting and gain only 24 tokens. Suppose that, due to moral preferences, the psychological cost from engaging in an activity that inflicts harm on another (the Type-Y subject) is 4, 8, 16, and 20 tokens, respectively. This means that the net gain from the activity for each respectively is $31 (35 – 4), $27 (35 – 8), $19 (35 – 16), and $15 (35 – 20). In such a case, X1 (who values acting at 31 tokens) and X2 (who values acting at 27 tokens) would act, but X3 and X4 would not (19, 15 < 24).

99 See supra Section I.A (analyzing the incentive scheme in Example 1 under a Liability regime).
Method allows us to observe subjects’ conditional strategies and test for the impact of the shared Liability regime directly.

**D. The Pilot**

Prior to running the full range of treatments, we conducted a pilot that allowed us to test the procedures and basic design of the experiments. Below we report briefly on the main results from the pilot sessions.

Sixty-four subjects participated in the pilot, which involved two Context-Free–Liability (CF–L) sessions and two Context-Free–No Liability (CF–NL) sessions, each with seven rounds. These treatments differed from the treatments used in the primary analysis\(^\text{100}\) in that pilot groups had only four members: three Type-X subjects and one Type-Y subject. A limitation of the pilot results is that the smaller group size and payoff incentives, combined with the use of the Strategy Method, created a dominant strategy for unconditional action in the Liability treatment.\(^\text{101}\) The treatments in the full experiment were modified to increase the number of Type-X subjects from three to four and adjust the monetary payoffs to avoid the creation of a dominant strategy in the Liability treatment.

**Figure 1: Average Unconditional Decisions by Liability Treatment in Pilot**

![Graph showing average unconditional decisions by liability treatment in pilot sessions.](image)

Although a detailed analysis of the pilot session is reported in Appendix I, it is worth noting two results here. For each treatment, Figure 1

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100 These treatments in the full experiment are reported below, *infra* Section I.E.

101 The pilot and the resulting dominant strategy are discussed in detail in Appendix I.
above shows the average unconditional decisions to act in each of the seven rounds. The vertical axis is the proportion of Type-X subjects who chose to act in a particular round. The horizontal axis is the round number. Notice that in every Context-Free round, the proportion of Type-X subjects choosing to act is higher in the CF–L (Liability) regime than in the CF–NL (No Liability) regime. In the CF–NL treatment, subjects were also much less willing to act alone than when joined by one or two others. These results are surprising given that the dominant strategy to act in CF–NL was more apparent than the dominant strategy to act in CF–L. The result may indicate Type-X subjects’ other-regarding preferences, such as kindness or guilt for the subject, who was left with zero tokens when one Type-X subject acted in CF–NL. The results are also consistent with the prediction that under the No Liability regime, an actor will be less inclined to engage in a tortious activity by herself but more inclined to do so with others. While not as consistent in the full experiment, we find similar evidence when the unconditional dominant strategy to act is removed. Importantly, the pilot’s results indicate that subjects understood the instructions and procedures and reacted consistently to the different liability regimes, as is also confirmed by the results in the full experiment.

E. The Full Experiment

The full experiment involved 200 subjects. We conducted three independent sessions for each of our treatments: the Context-Free treatment with No Liability (CF–NL) and Liability (CF–L), and the Vignette (Factories) with No Liability (V–NL) and Liability (V–L). As with the pilot, each session involved seven rounds with randomly rematched subjects for each round. A summary of the experiments is shown in Table 3.

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102 See infra Table A1-2.
103 See supra note 76; infra note 139 (discussing economic literature on individual preferences that incorporate aspects of others’ welfare); see also supra notes 94–98 and accompanying text (explaining the role of psychological constraints such as guilt or social disapproval on willingness to impose harm on others).
104 The risk and demographic controls are reported in Tables AII-4 and Table AII-5.
105 The average earning per subject was $18.20 per session, which included $1.38 from the risk elicitation task and $16.82 on average from the experiment.
TABLE 3: SUMMARY OF EXPERIMENTS

<table>
<thead>
<tr>
<th>Treatment</th>
<th>Group Composition</th>
<th>Tort Liability</th>
<th>Sessions</th>
<th>Subjects</th>
</tr>
</thead>
<tbody>
<tr>
<td>Context-Free Liability (CF–L)</td>
<td>4 Type-X 1 Type-Y</td>
<td>Liability</td>
<td>3</td>
<td>45</td>
</tr>
<tr>
<td>Context-Free No Liability (CF–NL)</td>
<td>4 Type-X 1 Type-Y</td>
<td>None</td>
<td>3</td>
<td>45</td>
</tr>
<tr>
<td>Vignette Liability (V–L)</td>
<td>4 Factories 1 Owner</td>
<td>Liability</td>
<td>3</td>
<td>50</td>
</tr>
<tr>
<td>Vignette No Liability (V–NL)</td>
<td>4 Factories 1 Owner</td>
<td>None</td>
<td>3</td>
<td>60</td>
</tr>
</tbody>
</table>

The summary statistics across sessions and rounds for the full range of treatments are shown in Table 4. We begin by comparing liability regimes in the Context-Free treatments, followed by an analysis of the Vignette treatments.

TABLE 4: DESCRIPTIVE STATISTICS ACROSS ALL ROUNDS

<table>
<thead>
<tr>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>Unconditional Decision to Act</td>
<td>68.25% (37.11%)</td>
<td>81.75% (27.87%)</td>
<td>85.36% (27.92%)</td>
<td>91.67% (11.35%)</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>0 other Type-X/Factory acting (0X)</td>
<td>15.08% (28.76%)</td>
<td>48.81% (41.74%)</td>
<td>12.50% (30.03%)</td>
<td>62.20% (45.22%)</td>
</tr>
<tr>
<td>1 other Type-X/Factory acting (1X)</td>
<td>29.76% (42.29%)</td>
<td>79.76% (32.46%)</td>
<td>23.57% (38.35%)</td>
<td>91.07% (20.55)</td>
</tr>
<tr>
<td>2 others Type-X/Factory acting (2X)</td>
<td>92.06% (15.06%)</td>
<td>87.70% (24.66%)</td>
<td>86.79% (27.14%)</td>
<td>86.90% (27.46%)</td>
</tr>
<tr>
<td>3 others Type-X/Factory acting (3X)</td>
<td>96.42% (11.00%)</td>
<td>88.89% (22.16%)</td>
<td>94.64% (16.71%)</td>
<td>87.80% (27.17%)</td>
</tr>
</tbody>
</table>

106 Percentages reflect the proportion of subjects—Type-Xs in the Context-Free treatments and Factories in the Vignettes—who decided to act, averaged across all seven rounds. Values in parentheses are standard deviations across actors’ (Type-X subjects’ and Factories’) average decisions to act. The difference in sample sizes for each treatment in Tables 3 and 4 is the forty innocent (nonacting) parties (Type-Y subjects and Owners) in the experiment.

107 In this case, N equals the number of Type-X or Factories in the treatments, as they are the decision-makers in groups. The total number of subjects in each treatment (as reported in Table 3) include the Type-Y subjects. All analysis based on subjects’ decisions to act does not include Type-Y or Owners.
1. The Context-Free Treatment: Tort Liability and Deterrence

The first hypothesis we seek to test with our experiments relates to the deterrent effect of tort law in cases involving group causation where liability is diluted as the number of liable tortfeasors increases.

a. Unconditional decisions to act

We begin by examining the unconditional decisions to act. On average, a smaller proportion of Type-X subjects chose to act under the Liability regime compared to No Liability—68.25\% versus 81.75\%—in the Context-Free treatment. Figure 2 shows the differences in averages across rounds in the Context-Free treatment. However, these apparent differences are not statistically significant.\textsuperscript{108} This result provides some support for our general hypothesis. The possibility that shared liability may dilute the costs of action appears to weaken the deterrent effect of tort liability to the point that its effect is statistically indistinguishable from willingness to act under a No Liability regime.

![Figure 2: Average Unconditional Decisions to Act by Round](image)

Formal regression analysis allows us to take advantage of the panel nature of the data. We use a multilevel logit panel regression to examine the likelihood a Type-X subject chooses to act (0 or 1). The primary independent variable is a dummy variable that captures the effect of liability introduced

\begin{flushright}
\textsuperscript{108} Using nonparametric Wilcoxon rank-sum tests, we reject the hypothesis that more subjects choose to act when they do not face liability ($p = 0.1456$, $n = 72$).
\end{flushright}
in the CF–L treatment relative to the omitted CF–NL treatment. The multilevel regression also controls for statistical dependence at the session and individual levels.\textsuperscript{109} Table 5 reports our logit-regression coefficients as odds ratios for unconditional decisions (Model 1) and for the various conditional decisions (Models 2–5). An odds ratio for CF–L is defined as the probability a subject in CF–L acts divided by the probability a subject in CF–NL acts. Thus, when the odds ratio is less than 1, the likelihood of acting when under Liability is relatively less than the likelihood of acting under No Liability. Conversely, an odds ratio greater than 1 would indicate that actors are more likely to act under Liability than under No Liability.

**Table 5: Mixed-effects logit regression—context-free treatment comparisons\textsuperscript{110}**

<table>
<thead>
<tr>
<th>Independent Variables</th>
<th>Model 1</th>
<th>Model 2 Cond.-0X</th>
<th>Model 3 Cond.-1X</th>
<th>Model 4 Cond.-2X</th>
<th>Model 5 Cond.-3X</th>
</tr>
</thead>
<tbody>
<tr>
<td>CF–L Dummy</td>
<td>0.25*</td>
<td>0.01***</td>
<td>&lt; 0.01***</td>
<td>1.41</td>
<td>6.01</td>
</tr>
<tr>
<td></td>
<td>(p = 0.066)</td>
<td>(p &lt; 0.001)</td>
<td>(p &lt; 0.001)</td>
<td>(p = 0.685)</td>
<td>(p = 0.160)</td>
</tr>
</tbody>
</table>

Starting with the unconditional decision to act, the coefficient for the CF–L dummy variable in Model 1 is less than 1 but only weakly significant at the 10\% level (\(p = 0.066\)). This implies that actors are less likely to act unconditionally under a Liability regime compared to a No Liability regime, although this evidence of a deterrent effect is not highly significant.

**b. Conditional decisions to act**

Turning to conditional willingness to act, subjects under a No Liability regime are more likely to act than those under a Liability regime, both when they act *alone* (48.81\% versus 15.08\%) and with only *one* other (79.76\% versus 29.76\%).\textsuperscript{111} These differences in percentages of subjects acting are highly statistically significant.\textsuperscript{112} This behavior is consistent with the theory. When facing liability, acting alone or with another is a losing strategy: it yields fewer benefits ($11 and $23, respectively) compared to not acting ($24). As predicted by the theory, tort liability deters in these circumstances.

\textsuperscript{109} See *infra* Appendix II for further details.

\textsuperscript{110} The full regression models are reported in Appendix II.

\textsuperscript{111} See *supra* Table 4.

\textsuperscript{112} Using nonparametric Wilcoxon rank-sum tests, differences in willingness to act with Liability versus No Liability conditional on 0 others and 1 other are statistically significant (each comparison \(p < 0.001, n = 72\)). Regression analysis also supports this conclusion as the CF–L coefficients in Models 2 and 3 of Table 5 are less than 1 and highly significant (\(p < 0.001\) in both comparisons).
By contrast, with a group of three or more actors—that is, when acting yields monetary gains under Liability—the proportion of Type-X subjects who were willing to engage in the tortious activity was actually higher with Liability than with No Liability (92.06% versus 87.70% for 2X and 96.42% versus 88.89% for 3X). An examination of the odds ratios in Table 5 provides additional support: conditional on two or three others, the odds ratio estimates switch from being less than 1 to being greater than 1—indicating that actors are more likely to behave tortiously under a Liability regime. While these differences across liability regimes are not statistically significant, the results show that once joined by enough actors, the dilution of liability renders any deterrent effect of tort law indistinguishable from a regime with No Liability.

The results are consistent with our first hypothesis—that group causation theories erode the deterrent effect of liability regimes. They are also consistent with the tortfest dynamic we observe. As the number of actors increases, liability is further diluted, the individual monetary benefits increase, and, consequently, more are willing to engage in the tortious behavior. In the Context-Free–Liability (CF–L) treatment, conditional on two others acting, when acting becomes profitable, liability loses its deterrent effect. Conditional on three others acting, as the gain increases, our point estimates indicate actors may be even more likely to engage in the tortious behavior compared to a No Liability regime. This may seem odd at first. But a number of reasons could support the pattern: First, under a Liability regime, the victim is compensated. Actors may thus experience a reduction in their moral cost compared to No Liability. Second, as we show below, we observe that fairness-rating levels increase with the number of tortfeasors, thus suggesting another decrease in nonpecuniary costs. Third, the increase in the number of actors increases the individual monetary payoff of each actor. Finally, as the individual payoffs become positive, so do the net social gains.

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113 See supra Table 4.

114 The individual odds-ratio estimates are not statistically different from 1 due to large standard errors. P-values are found in Table 5 for the individual estimates. The lack of significance here only illustrates the absence of a deterrent effect for tort law at these higher participation levels.

115 Using nonparametric Wilcoxon rank-sum tests, the difference in willingness to act with Liability versus No Liability is not statistically significant conditional on two others and is weakly significant conditional on three others (\( p = 0.931 \) conditional on two others acting and \( p = 0.105 \) conditional on three others, \( n = 72 \)). However, multilevel panel regressions find treatment differences are not statistically different for both comparisons. The coefficients in Models 4 and 5 of Table 5 are greater than 1 but not statistically significant (\( p = 0.685 \) conditional on two others acting; \( p = 0.160 \) conditional on three others acting).

116 See infra Section II.E.2.
Thus, actors may be able to view their actions and the ensuing harm to the compensated victim as necessary for the greater good.\textsuperscript{117}

2. Context-Free Treatment: Tortfest and Causation

\textit{a. Tortfest}

Our second and third hypotheses relate to the potential for tort law itself to induce multiple subjects to join in the tort—the tortfest dynamic. In part, this dynamic arises because but-for causation is operative, even though, according to tort law jurisprudence and scholarship, it does not function in concurring-causes cases.\textsuperscript{118}

By analyzing the conditional decisions made in the Strategy Method, we provide support for the tortfest dynamic. Table 4 shows that the proportion of subjects willing to commit a tort under the Liability treatment (when averaged across all decision rounds) increases as the number of others committing the tort increases. Conditional on zero and one other Type-X subject committing a tort, the proportion of Type-X subjects choosing to act increases from 15.08% to 29.76%. While this is higher than the theoretical prediction of 0%, the tortfest and but-for predictions are extremely stark when examining the increase in willingness to commit a tort conditional on two and three others also committing the tort: 92.06% and 96.42%, respectively. It is clear that subjects understand the incentives created by the Liability regime and would like to engage in group “wrongdoing” if a large enough number of participants join them.

\textsuperscript{117} In this sense, acting is similar to a Pareto Public Good that improves overall welfare in a group, even though it may actually hurt some group members. Sagi Dekel, Sven Fischer, & Ro’i Zultan, \textit{Potential Pareto Public Goods}, 146 J. PUB. ECON. 87, 87–88 (2017) (describing and providing examples of Pareto Public Goods and testing voluntary contributions to such goods).

\textsuperscript{118} See \textit{supra} notes 5, 10–11; \textit{supra} Section I.B.
Figure 3: Conditional Decisions to Act Under Liability and No Liability Treatments by Round

The top graph in Figure 3 above shows that under the Liability treatments, we observe a monotonic increase in the average proportion of Type-X subjects acting as we move from decisions conditional on zero to one, two, and three other Type-X subjects participating. Each of these jumps in Type-X’s conditional willingness to participate is statistically significant. This result contradicts the assumption that extending liability

\[ p = 0.022; \text{to} \ 92.06\%, \ p < 0.001; \text{and to} \ 96.42\%, \ p = 0.055, \ n = 36). \]

Panel regression analysis comparing individuals’ conditional decisions in all decision rounds controlling for individual-level and session-level dependencies is not possible. An alternative robustness check is discussed in Appendix II.
will always deter actors in a group, instead showing it can produce a dynamic in which subjects become more willing to act as they are joined by others. In the treatments with No Liability (the bottom graph in Figure 3 above), we see a different pattern. We still observe a similar increase in conditional willingness to participate in tortious activity as the number of participants grows: from 48.81% (conditional on no one else acting) to 79.76% (conditional on another acting), rising to 87.70% (conditional on two others acting), and to 88.89% (conditional on three others acting). However, unlike the Liability treatment, the largest, statistically significant jump is between willingness to act conditional on no one else acting (that is, 0X) and willingness to act conditional on one other person acting (1X). The differences in willingness to act after that (that is, between 1X, 2X, and 3X) are smaller and not significant at the 5% level.\textsuperscript{120} Although Type-Xs in the No Liability treatments are clearly reluctant to act alone, even without facing liability, this hesitation generally disappears once they are joined by even one other Type-X. From that point on, the participation of additional Type-Xs is not systematically influential in the same way we see under the Liability regime.

It is in our Liability treatment that we see a marked difference in the effect of an additional tortfeasor when moving from 1X to 2X, where the additional actor makes the tort become profitable. Under shared liability, as the tortfest theory predicts, the willingness of our Type-X to participate also increases systematically with the number of actors beyond this point. These differences from No Liability provide strong support for the tortfest effect and its weakening of deterrence.

\textit{b. But-for causation}

We also find strong evidence suggesting that the but-for test can operate in cases involving group causation such as concurrent causes, thus supporting our third hypothesis. Under the Liability regime, we observe how but-for causation influences the results by examining the Type-Xs’ conditional decisions. As discussed above and shown in Figure 3, in general Type-X subjects are unwilling to act unless joined by at least two others. Below this threshold, at most 30% of Type-Xs would act on average. This is consistent with the theoretical prediction discussed in Part II. With only one or two Type-X subjects acting, we observe very little willingness to act. But

\textsuperscript{120} Using nonparametric Wilcoxon signed-rank tests comparing average conditional willingness to act across all rounds, the increase in Type-X’s conditional willingness to participate is statistically significant between conditional on 0X and 1X ($p < 0.001$, $n = 36$). The further increase between conditional on 1X and 2X is weakly significant (79.76\% to 87.70\%, $p = 0.089$), and the increase between conditional on 2X and 3X is not significant (87.70\% to 92.06\%, $p < 0.646$, $n = 36$). An alternative robustness check using Round 1 data only is discussed in Appendix II.
with two or more Type-X subjects participating, actors seem to be eager to join the tortfest because of the dilution of liability. Graphically, the top graph in Figure 3 above illustrates the but-for dynamic, by creating a clear, very large, and statistically significant jump from acting conditional upon one other subject acting as well (1X) to acting conditionally upon two other subjects acting (2X)—more than a 62% increase in willingness to act (from 29.76% to 92.06%). We observe the same, equally strong effect in the Vignette (23.57% to 86.79%) and the pilot (12.50% to 95.24%).

The increased willingness to act conditional on two or three others acting together is consistent with the financial-incentive effect for a tortfest created by shared liability. This effect and our results provide additional evidence that a tortfest would not develop but for participation by others. The increased willingness to join—conditional on two or three others—may also reflect a decrease in psychological (nonpecuniary) costs. For example, participants may believe that a larger number of tortfeasors indicates less moral culpability from acting. In such a case, one would expect that more actors would be willing to gradually join the tortfest as the number of actors increases. We indeed find strong evidence for this fairness hypothesis.

The importance of but-for causation in our results is clear. The number of Type-X actors who are willing to participate unconditionally in the tortious act depends critically on their expectations about what others will choose. The relatively high proportion of Type-X who act unconditionally under the Liability regimes (68.25% in CF–L and 85.36% in V–L) reflects actors’ general belief that they will in fact be joined by others. But for the belief that a third or a fourth person will also engage in the tortious activity, their conditional decisions indicate they would not act, and the harm to the Type-Y would not occur.

3. Context: Reexamining Tortfest and Deterrence

The fourth hypothesis relates to the way the context of the tort setting interacts with tortfests and deterrence. To test the effect of context, we compare decisions to act in the neutral Context-Free frame to the Factory Vignette. We predict that the willingness to commit the tort will be higher in our Factory Vignette than in the Context-Free frame, as the Factory Vignette makes the social benefit from the tort (i.e., producing cement) more salient compared to the Context-Free frame.

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121 Note that in the pilot, the increase was observed between conditional on 0 and conditional on 1 other subject acting. Groups only had three Type-X subjects and the payoffs were parameterized differently. See infra Appendix I: Pilot Results.
122 See infra Section II.E.4.
123 See Alekseev et al., supra note 71 (examining the effect of contextualized instructions in economic experiments); Raymond & Cason, supra note 81 (same).
a. Unconditional decisions to act

Table 4 reports that unconditional average proportions of willingness to act are higher in the Vignettes than in the Context-Free treatments. This is true under Liability (V–L 85.36% versus CF–L 68.25%) and under No Liability (V–NL 91.67% versus CF–NL 81.75%). In other words, regardless of the liability regime, more actors are willing to unconditionally commit a tort in the contextualized treatment.

**Figure 4: Unconditional Decisions to Act Across Liability Regimes: Vignette vs. Context-Free Framings**

The differences are also clearly observed in Figure 4, which displays the proportion of subjects (Type-Xs and Factories) acting unconditionally in each round. In Figure 4, the graph on the top displays the data for acting unconditionally in the Liability treatments (CF–L and V–L). The graph on
the bottom is corresponding data for the No Liability treatments (CF–NL and V–NL). It is easy to see that the proportion of subjects (Type-Xs and Factories) acting is higher in the Vignette treatments for both liability treatments in every round. This contextual difference is highly statistically significant with Liability and weakly significant with No Liability.\textsuperscript{124}

Recall that the first hypothesis predicts reduced deterrence of tort law as a result of imposing liability on multiple actors, which at some point may even lead to higher participation in tortious activity than when actors face no liability.\textsuperscript{125} The results from our contextualized unconditional decisions support this reduced-deterrence effect. The average proportion of subjects acting unconditionally in V–L (85.36%) is not significantly different than the average proportion in V–NL (91.67%).\textsuperscript{126} The finding that liability does not significantly influence unconditional willingness to act is robust across both our nonparametric tests and regression results in the Vignette setting. If anything, context appears to strengthen this reduced-deterrence effect.

<table>
<thead>
<tr>
<th>Independent Variables</th>
<th>Model 6</th>
<th>Model 7</th>
<th>Model 8</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Odds Ratio</td>
<td>Odds Ratio</td>
<td>Odds Ratio</td>
</tr>
<tr>
<td>V–L Dummy</td>
<td>8.78***</td>
<td>—</td>
<td>0.65</td>
</tr>
<tr>
<td>( p = 0.008 )</td>
<td></td>
<td></td>
<td>( p = 0.453 )</td>
</tr>
<tr>
<td>V–NL Dummy</td>
<td>—</td>
<td>2.54*</td>
<td>—</td>
</tr>
<tr>
<td>( p = 0.075 )</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

\textit{b. Conditional decisions to act}

Hypotheses 2 and 3 concern the tortfest dynamic and but-for causation. We examine the effect of the more realistic context in our Vignettes on these hypotheses by again looking to the conditional decisions that subjects made using the Strategy Method.

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\textsuperscript{124} Nonparametric rank-sum tests report that the median difference in willingness to act across contextualized instructions is significant with Liability (\( p = 0.025, n = 76 \)) but not with No Liability (\( p = 0.262, n = 84 \)). However, formal multilevel logit regression analysis reported in Models 6 and 7 of Table 6 finds that differences between Neutral and Vignette treatments are significant at the 1% level with Liability (Model 6, \( p = 0.008 \)) and at the 10% level with No Liability (Model 7, \( p = 0.075 \)).

\textsuperscript{125} \textit{Supra} Section II.C ("We hypothesize that the imposition of liability on multiple actors will incentivize an actor who would alone engage in a harmless activity to switch to a tortious activity if liability is imposed on a large enough number of actors.").

\textsuperscript{126} This is verified in a nonparametric rank-sum test (\( p = 0.934, n = 88 \)) and multilevel logit regression analysis reported in Model 8 of Table 6 (\( p = 0.453 \)).

\textsuperscript{127} The full regression models are reported in Appendix II.
We begin by comparing conditional decisions with Liability under the Vignette and Context-Free treatments (i.e., CF–L versus V–L). Our results here appear to be robust to context. The average proportions in Table 4 for CF–L and V–L are all very similar and not statistically different: conditional on zero others (15.08% versus 12.50%); conditional on one (29.76% versus 23.57%); conditional on two (92.06% versus 86.79%); and conditional on three (96.42% versus 94.64%). These results are confirmed by regression analysis reported in Table 7 below. The omitted dummy variable in each Model is a variable for CF–L. While the odds ratios are all less than 1, indicating a lower likelihood of acting in the V–L condition relative to the Context-Free frame, these differences are nowhere near statistical significance. These results are reassuring in that they suggest the tortfest dynamic and but-for causation effects under Liability treatment are robust to presentation in a more realistic context. At the level of conditional decisions, the more realistic Vignette treatment does not appear to lead to individuals being more willing to act, as we had hypothesized. However, the statistically significant difference in unconditional willingness to act in the Liability treatment perhaps suggests it is easier for actors (i.e., subjects assigned as Type-X or Factories) to implicitly predict others’ participation in the tortfest that leads to their own higher willingness to act.

<table>
<thead>
<tr>
<th>Independent Variables</th>
<th>Model 9 Cond.-0X Odds Ratio</th>
<th>Model 10 Cond.-1X Odds Ratio</th>
<th>Model 11 Cond.-2X Odds Ratio</th>
<th>Model 12 Cond.-3X Odds Ratio</th>
</tr>
</thead>
<tbody>
<tr>
<td>V–NL Dummy</td>
<td>0.81</td>
<td>0.52</td>
<td>0.60</td>
<td>0.74</td>
</tr>
<tr>
<td>Dummy</td>
<td>p = 0.850</td>
<td>p = 0.632</td>
<td>p = 0.547</td>
<td>p = 0.815</td>
</tr>
</tbody>
</table>

In addition to testing the impact of context within the Liability treatment, we also assess the robustness of our tortfest dynamic and but-for causation results across liability treatments (e.g., V–L versus V–NL). Figure 5 visually confirms that the same patterns observed in conditional decision-making in the Context-Free treatments are also present in the more realistic Vignette treatments. Figure 5’s top graph illustrates the tortfest dynamic in

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128 Nonparametric rank-sum tests (n = 84) and multilevel panel logit regressions reported in Models 9–12 of Table 7 confirm that the differences between conditional decisions in Context-Free and Vignette treatments with Liability are not statistically significant (p > 0.54 for each comparison).

129 The full regression models are reported in Appendix II.

130 The conditional decisions to act in the Context-Free treatments are described supra in Figure 3.
V–L, with a substantial and highly significant increase between acting conditional on one other to conditional on two others acting.\footnote{Comparing average conditional willingness to act across all rounds, Wilcoxon signed rank, \( p < 0.001, n = 40 \). An additional robustness check is examined in Appendix II.}

**Figure 5: Average Conditional Decisions to Act in Vignette: Liability vs. No Liability by Round**

Figure 5’s bottom graph illustrates the proportion of Factories acting conditional on others acting under a No Liability regime. Again, as in CF–NL, in V–NL there is a significant increase in willingness to act between acting alone and acting conditional on one other joining the activity (0X to
In other words, it appears that in the contextualized No Liability regime, a substantial proportion of subjects are reluctant to act on their own to secure a gain at another’s cost. However, once joined by at least one other person, subjects are significantly more willing to commit the tort (91.07% versus 62.20%). Willingness to act alone appears higher in the Vignette–No Liability frame than the corresponding figure in the Context-Free frame (V–NL versus CF–NL conditional 0X), which perhaps suggests that the benefits are more salient in the contextualized frame.

4. Tort Liability and Fairness

Our final hypotheses relate to the way in which tort liability interacts with the perceived fairness of an individual’s decision to act. Our experimental design allows us to examine the ways in which the private costs and benefits of actors are balanced against fairness concerns for the potential victim. We do so by comparing unconditional and conditional decisions to act across our liability conditions. We also use subjects’ responses to the fairness questions. We hypothesize that tort liability can potentially blunt fairness concerns and may even create a tortfest dynamic in fairness perceptions, rather than signaling moral culpability.

It is unclear ex ante how a liability regime may interact with subjects’ concerns about fairness. Prior studies suggest that there may be a “norm-activation” impact that helps deter activity that imposes harms on others or a deterrence aspect that arises through an expressive function of tort liability in declaring and denouncing behavior from a social perspective. However, a more general literature supports an alternative: that subjects may treat the liability obligation as the “price” to engage in the activity. Under this view, liability more readily allows actors to weigh the “cost” of acting

132 Comparing average conditional willingness to act across all rounds, Wilcoxon signed rank, \( p < 0.001, n = 48 \). An additional robustness check is examined in Appendix II.

133 Comparing average willingness to act conditional upon 0X across all rounds, Wilcoxon signed rank, \( p = 0.104, n = 84 \). Note that this difference is at the margin of conventional statistical significance.


135 See Deffains et al., supra note 62, at 2 ("[T]he introduction of legal liability has a norm-activation effect, in the sense that individuals then put greater weight on losses caused to others, which complements the threat of legal sanctions.").

136 See Eisenberg & Engel, Unpacking Negligence Liability, supra note 45, at 134 (noting deterrent effects from courts that “make[] social expectations explicit” and “blame” defendants who fail to meet them as distinct from incentives created by pecuniary liability).

137 See supra notes 94–95 and accompanying text.
against their own private benefits from acting. In the case of shared liability and the potential for a tortfest dynamic, an additional possibility is that as more individuals join the tortfest, it may be perceived as less harmful and more socially acceptable. The existence of the liability rule may help to focus subjects’ attention on the shared benefits and costs of the group and lower psychological costs that may be associated with distributive aspects of the activity.

a. Observed behavior

One of the first points to note from our experimental results is that subjects are generally reluctant to act at the expense of an innocent victim if they are the only Type-X to benefit from the activity. This is perhaps most evident in the average conditional willingness to act in the scenario with No Liability. In the Context-Free frame, only 48% of Type-X subjects indicate willingness to act alone, although the monetary payoff—35 tokens versus 24 tokens—is significant. In the Vignette frame, more subjects are willing to act alone, but 37% of subjects would not act unless accompanied by others, even though acting under a No Liability regime promises a gain at no monetary costs.

However, once subjects are no longer acting alone, the pattern across liability treatments begins to diverge. In the No Liability regime, Type-X individuals are much more willing to act to gain the benefit at Type-Y’s expense. Conditional on one actor joining, the willingness to commit a tort jumps to 79% in the Context-Free frame and 91% in the Vignette. These proportions are far higher than in the treatments where Type-X’s face Liability: conditional on one other acting, 29% are willing to act in the


139 A large literature confirms that subjects often depart from behavior predicted by self-interested monetary payoffs in experimental settings and, instead, appear to behave in ways consistent with a degree of regard for the welfare of others and an aversion to inequality. See, e.g., Gary Charness & Matthew Rabin, *Understanding Social Preferences with Simple Tests*, 117 Q.J. ECON. 817, 817 (2002) (testing the motivations of participants in economic experimental games and finding a primary concern for social welfare); John A. List, *On the Interpretation of Giving in Dictator Games*, 115 J. POL. ECON. 482, 482 (2007) (providing a comprehensive study of results from dictator games and finding persistent result that participants do not choose the most selfish outcome); JOSEPH HENRICH, ROBERT BOYD, SAMUEL BOWLES, COLIN CAMERER, ERNST FEHR & HERBERT GINTIS, *Foundations of Human Sociality: Economic Experiments and Ethnographic Evidence from Fifteen Small-Scale Societies* 8 (2004) (describing experimental tests carried out in a large, cross-cultural study finding no society in which behavior was consistent with pure self-interest).

140 See * supra* Table 4.
Context-Free frame with Liability and 23% in the Vignette. These differences are strongly statistically significant.

However, once two other subjects are acting, the proportions of those who are willing to act on average converge across liability treatments. In the Context-Free frame, 92% act with Liability versus 87% with No Liability. In the Vignette frame, the corresponding proportions are 86% for both treatments. There is no statistically significant difference across liability regimes once three actors are involved. Interestingly, conditional on three actors joining, the share of actors willing to act becomes higher in the Liability regime, rising in the CF–L to 96%, compared to 88% in the CF–NL, and in the V–L 94%, compared to 87% in the V–NL. By contrast, there is virtually no change in the No Liability conditions: 88% in the CF–NL compared to 87% in V–NL.

These differences at the higher levels of expected actors’ participation imply that tort liability may be incentivizing participation in the harmful activity by reducing the monetary cost to the actors and also by mitigating fairness concerns. Behavior in the No Liability treatments appears consistent with subjects generally being willing to act for private gain with a loss to the victim—so long as they are not solely responsible. However, actors seem perhaps less willing to participate in some situations where they benefit and another is harmed without redress than in the case where liability provides compensation to the victim. Compensation to the victim potentially helps the actors focus on the collective benefits and dilutes moral culpability from acting as more participants join in the tort.

See supra Table 4.

Testing for equivalence across liability treatments in the Context-Free frame, we reject with \( p < 0.0001, n = 72 \), and in the Vignette frame \( p < 0.0001, n = 88 \).

See supra Table 4. While the higher average willingness to act in the Liability treatments is not generally significant, the results approach significance in the Context-Free frame. Testing for equivalence between CF–L and CF–NL conditional on the others acting can be rejected, but only weakly, as \( p = 0.1050, n = 72 \).

This is consistent with finding conduct more morally and legally culpable when an actor intentionally inflicts harm on a victim as opposed to exposing her to a risk of harm. See, for example, the discussion in Eisenberg & Engel, Unpacking Negligence Liability, supra note 45, at 117, stating, “In light of prevalent moral intuitions, knowingly imposing harm on an innocent outsider is not the same as exposing this outsider to the risk of suffering harm.” It also reflects theories that ascribe greater blameworthiness as individuals become more directly and consciously responsible for causing harm. See Alicke, supra note 37, at 563 (assessing blameworthiness based on various dimensions of control; control is highest “when the actor purposely (volitional behavior control) and with foresight (volitional outcome control) causes harmful consequences (causal control)”).

Eisenberg & Engel, Unpacking Negligence Liability, supra note 45, at 137 (discussing their findings suggesting that provision of compensation “crowded out” moral concerns related to tortiously harming a victim).
b. Fairness attitudes

These intuitions appear to be consistent with results from the fairness questions. In this set of questions administered following the seventh round in each session, we asked subjects to rate the fairness when one, two, three, and four actors act. Subjects were asked to assess overall fairness and also asked to assess the fairness to the those who acted, the victim, and for all five subjects—i.e., the four Type-Xs and the Type-Y in the Context-Free frame and the four Factories and Owner in the Vignette. We used a scale from 0 to 5, 5 five being “completely fair,” allowing us to assess fairness both within and across liability treatments.
Figure 6: Response to Fairness Questions: Fairness to Actors, Victim, and Entire Group
Figure 6 displays the average fairness rating to the actors, the victim, and the entire group. Within each of the three graphs, there are average fairness points for each possible scenario: one subject acting, two subjects, three subjects, and all four subjects in the group acting. To increase the questionnaire’s statistical power, data are pooled by liability regime. For instance, the blue line in each of Figure 6’s graphs illustrate the average fairness ratings trend for subjects in CF–NL and V–NL together. Importantly, these average fairness ratings only contain responses from subjects assigned as actors, as we are particularly interested in the responses of subjects who made the decision whether or not to act.

The top graph in Figure 6 displays the average fairness rating to the actors. With No Liability, the fairness ratings are relatively stable as the number of actors increases. By contrast, with Liability there is a clear upward trend. Comparing fairness with one acting to fairness with four acting, the difference is statistically significant at the 5% level. The trend is so pronounced that, while the average fairness rating when one subject acts is significantly higher with No Liability, the average fairness rating when four subjects act is significantly higher with Liability. This marked rise in the perceived fairness of the activity under the Liability regime is consistent with tort liability acting to facilitate rather than deter the activity (by lowering the psychological costs that may be associated with acting alone to cause harm or benefiting at the expense of a single victim).

The middle graph in Figure 6 displays the average fairness rating to the victim. The clearest observation from this graph is the large difference between the liability regimes in every scenario. Comparing average fairness ratings across liability regimes, fairness is higher with Liability than with No Liability and is significant in every case. Also note that, since 2.5 is the middle of the fairness rating range, actors always felt that acting was unfair to the victim with No Liability, regardless of the number of acting subjects. By contrast, Liability works the other way. In the Liability regime, actors always felt, on average, that engaging in the harmful activity was fair to the

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146 This is a nonparametric signed-rank test, \( p = 0.028, n = 6 \). Each session is treated as an independent observation. The small sample sizes for the tests of fairness are the conservative approach to controlling for subjects’ interactions during the experiment. Recall that with Strangers Matching during the experiment, subjects were randomly rematched with others in the same session. This leads to alternative approaches to control for such independence. The least conservative approach would be to treat each subject as an independent observation. Appendix II provides a discussion of these approaches.

147 This is a nonparametric rank sum, \( p = 0.025 \) and 0.005, respectively, \( n = 12 \).

148 See supra notes 94–98, 135–139 and accompanying text for discussion of the theories behind psychological costs.

149 This is a nonparametric rank sum, \( p < 0.004 \) in each case, \( n = 12 \).
victim. However, comparing the case with one subject acting to the case with four acting, the fairness ratings significantly increase with No Liability but not with Liability.\footnote{This is a nonparametric signed rank, $p = 0.046$ and 0.600, respectively, $n = 6$.}

The bottom graph in Figure 6 displays the average fairness ratings to the entire group—a proxy for “society.” This graph presents a clear pattern of fairness ratings: in each liability regime, fairness significantly increases as the number of actors increases.\footnote{This is a nonparametric signed rank, $p = 0.028$ for each liability regime, $n = 6$.} The other apparent pattern is that, in every case, average fairness ratings are greater with Liability than with No Liability regimes.

To summarize, we observe the tortfest dynamic in the fairness ratings to the entire group—a result we did not expect. Although there is a level effect when it comes to the liability regime, fairness ratings to the entire group are consistently and significantly higher with Liability than with No Liability. It is also interesting to point out that with No Liability, fairness ratings to the entire group start out slightly unfair (2.33) and end in the fair portion of the scale (3.35).

5. **Regression Results with Controls**

In the previous Sections we reported multilevel logit regression analyses examining treatment differences in individuals’ decisions to act or not. In these regressions, the independent variables were treatment dummy variables and a time trend, which controlled for the round of the session in which the observations were generated. In addition to these regressions, we ran several models using additional demographic and other controls. These additional controls include: (1) measures of subjects’ risk proclivity, (2) the number of previous economics experiments that the subjects have participated in, (3) a male dummy variable,\footnote{Subjects could identify as male, female, transgender, or other.} (4) a dummy variable for a subject with a business major, (5) a dummy variable for a subject who is white, (6) a dummy variable for a subject who self-reports to be a believer in any supernatural being, (7) dummy variables for whether a subject is a Democrat or Republican (independent is the omitted dummy), and (8) a subject’s political-party intensity (0 to 5, 5 being very intense). The regression models with these additional controls pooled individuals’ observations across liability treatments and used dummies for contextualized instructions.\footnote{The individual data are panels, with seven observations for each individual subject for each round in the session.} As mentioned in Appendix II, we incorporated random
effects at the individual and session levels. We also included a round variable to control for any time trend (e.g., learning).\footnote{These regressions are reported in Appendix II.}

Regarding the differences between Context-Free and Vignette treatments, the results reported earlier are qualitatively confirmed in regressions with these additional controls. We find a significantly higher likelihood to act unconditionally in Vignette treatments than in Context-Free treatments with both Liability and No Liability regimes.\footnote{With a Liability regime, the likelihood of acting is significantly higher in the Vignette treatment than the Context-Free treatment (p = 0.034). Similarly, with a No Liability regime, the likelihood of acting is higher in the Vignette treatment (p = 0.020).} Decisions to act conditional on zero or one other actor are significantly more likely to occur in the Vignette treatment with No Liability. However, the Vignette treatment does not lead to statistically different likelihoods to act conditional on two or three other subjects acting. With Liability, the likelihood of acting does not change with the Vignette compared with the Context-Free instructions in any of the four conditional decisions.

Although we include a variety of controls, in general, these are not systematically significant in our regressions. This is somewhat reassuring, as it suggests that our results are not driven by particular characteristics of our sample.

\textbf{Conclusion}

The use of group causation theories such as substantial factor, concerted action, and alternative liability is assumed to deter potential injurers and prevent harm to victims. However, to date, the discourse regarding the deterrent effect of group causation theories has been based on intuition and conjecture. This discourse has not been attentive enough to the potential that liability will incentivize tortious activity. It has also lacked the critical support that empirical evidence provides. In this Article, we undertake to fill this gap by testing, for the first time, the effect of causation law on group wrongdoing. Using an incentivized experiment in a lab setting, we tested different liability regimes in both contextualized and neutral framings. Our analysis of over 1,200 data points from 200 subjects confirms that group causation doctrines, such as substantial factor and concerted action, can encourage, rather than discourage, actors to engage in tortious activities.

In the Liability regime, we observed a clear tortfest dynamic. In the Context-Free framing, the willingness to engage in the tortious activity increased with the number of tortfeasors from 15% (acting alone), to 29% (joined by one other), to 92% (with two others), to 96% (with three others). We observed the same dynamic in the Vignette framing: willingness to
behave tortiously increased with the number of actors from 12% to 23% to 86% to 94%. These jumps were statistically significant at the 1% level. In sum, despite—in fact, because of—the imposition of liability, the tortious activity became more beneficial and the pressure to join the activity grew stronger as more actors joined.

An analysis of the unconditional decisions to act in the liability regimes reveals that actors were also willing to take a “leap of faith.” Although acting tortiously was worthwhile only if an acting subject was joined by two others, the high willingness to act unconditionally—68% in the Context-Free framing and 85% in the Vignette framing—implies that actors believed or expected that others would choose to behave tortiously as well. The results also reveal that actors were more willing to behave tortiously, and unconditionally so, in the polluting-factories Vignette than in the Context-Free framing. One possible explanation is that in the contextualized framing, the costs to the victim and the benefits from the tortious activity are more salient. Consequently, actors may more easily see that they can reap a private gain and increase total welfare at a cost to a compensated victim. The result is similar to the one observed in a potential Pareto Public Good experiment.\(^{156}\)

In addition to evidence of the financial incentives from a tortfest, our results show that tort law can also blunt fairness concerns about acting in a way that harms others. Under the liability regimes, participants experienced the analog to a tortfest in their perceptions of fairness to the group. Just like the willingness to engage in a tortious activity, the level of perceived fairness to the group increased monotonically with the number of acting subjects.\(^{157}\) In other words, participants viewed the tortfest as more fair as the number of tortfeasors increased.

Moreover, under both framings, more actors were willing to engage in the tortious activity under Liability as compared to No Liability, conditional on being a part of a group of four tortfeasors (96% versus 88% in the Context-Free frame and 94% versus 87% in the Vignettes). While some of the results are only weakly significant, they suggest that the imposition of liability did not discourage actors and may have even done the opposite. This result initially may seem counterintuitive, but we offer a number of possible

\(^{156}\) See Dekel et al., supra note 117, at 87–89 (finding that when the public good harms a minority, group members reduce their contributions unless group members can communicate with each and send monetary transfer to the harmed minority).

\(^{157}\) This is also apparent in the consistently high proportion of individuals who are conditionally unwilling to act alone to cause harm to the Type-Y subjects, even when they do not face liability. The social dimension to fairness perceptions is also apparent in our No Liability treatment. Fairness perceptions in this variant also increase as more participants join the harmful activity. This is important because it suggests that there may even be potential for a synergistic effect: shared liability creates financial incentives for tortfests that may also transform perceptions of fairness.
explanations: A Liability regime may crowd out social norms and impose what may be perceived as a “price” that is discounted with the number of acting subjects. This discounted price may also suggest a low moral culpability. The tortious activity may even be justified by a “mob mentality”—an everyone-does-it rationale. It could also be viewed as morally acceptable when the activity produces a product from which society benefits, notwithstanding the harm to the (compensated) individual victim.

We also find clear evidence supporting the but-for test’s applicability in the Liability regime. Very few actors chose to act by themselves or with another (15% and 29% in the Context-Free treatment and 12% and 23% in the Vignette). However, consistent with the theory, the willingness to engage in the tortious activity jumped significantly conditional on two others joining them—when the tortious activity became profitable—increasing to 92% in the Context-Free framing and 86% in the Vignette. This willingness reached a whopping 96% in the Context-Free framing and 94% in the Vignette group of four actors. The jump from acting with one other to acting with two others (29% to 92% in the Context-Free frame and 23% to 86% in the Vignette) is substantial and statistically significant at the 1% level. It supports the conclusion that, consistent with the theory, the large majority of subjects (71% in the Context-Free frame and 77% in the Vignette frame) would not engage in the tortious activity but for two or more joining them. Formal regression analysis supports these results.

Our study also makes several important contributions to the prior economic and empirical literature. Previous studies focused on liability’s effect on a single tortfeasor, but they could not observe decisions at the individual level. Previous studies also could not test how imposing liability on many actors impacts the members of the group. By contrast, our unique design allowed us to directly test the effect of causation law in cases involving multiple actors. We were also able to observe individuals’ decisions, including whether and to what extent they are dependent on the actions of others.

Still, a key question for any experimental study in a laboratory setting is how relevant it may be in the “real world.”158 Does our experimental study

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158 A natural question when using laboratory experiments with student samples is whether behavior observed in student samples is generalizable to the behavior of the population. Professor Vernon Smith argues that a major benefit of controlled experiments is their internal validity based on “induced” institutions and environments. See Vernon L. Smith, Economics in the Laboratory, 8 J. ECON. PERSPS. 113, 113 (1994). However, Professors Steven Levitt and John List question the external validity of experiments to settings in the field. Steven D. Levitt & John A. List, What Do Laboratory Experiments Measuring Social Preferences Reveal About the Real World?, 21 J. ECON. PERSPS. 153, 154 (2007). Professor Colin Camerer critiques the arguments in Levitt and List. Colin F. Camerer, The Promise and
reflect insights important in the practical world of tort law? While there are some features of our experimental design that may suggest caution in extending our findings, we believe that the basic incentive structure underlying our experiment potentially applies in numerous settings where causation law creates the potential for shared liability. One such common setting is medical malpractice cases where multiple actors may be responsible for an indivisible harm to the victim. For example, doctors and nurses performing a surgery using tools manufactured and produced by third parties might cause an indivisible harm to the patient. Other settings, more closely aligned with our experiment, could include environmental harms, such as decisions about whether or not to properly dispose of potential contaminants. For example, multiple industry participants may contribute to groundwater contamination. Extending liability to catch all the contributing parties may instead lead to a dilution of costs that can incentivize the behavior.  

We end with a call for additional research. The use of an experimental methodology necessarily imposes some limitations in order to preserve internal validity and the ability to draw causal conclusions. In our experiment, we have focused on a setting with a loss that can be clearly monetized. However, many tort disputes involve personal injuries, not just injuries to property. In these settings, victims’ losses are more difficult to quantify, and actors may also experience higher nonpecuniary costs in addition to tort liability, whether moral, psychological, or social-norm based. Accordingly, some care would be required in extending our results to the context of personal harms. This area would be an important future extension of our current work. However, our experiment provides evidence that shared liability can support but-for causation and create incentives for tortfeasors that seriously weaken the disincentive effects of tort law. Our results caution against extending group liability based on the assumption it will always advance deterrence.

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APPENDIX I: PILOT RESULTS

The pilot involved two CF–L sessions and two CF–NL sessions. Each session had seven rounds and sixteen subjects, for a total of sixty-four subjects, as summarized in Table AI-1.

<table>
<thead>
<tr>
<th>Treatment</th>
<th>Group Composition</th>
<th>Type-X Liability</th>
<th>Sessions</th>
<th>Subjects</th>
</tr>
</thead>
<tbody>
<tr>
<td>Liability</td>
<td>3 Type-X</td>
<td>Liable</td>
<td>2</td>
<td>32</td>
</tr>
<tr>
<td></td>
<td>1 Type-Y</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No Liability</td>
<td>3 Type-X</td>
<td>None</td>
<td>2</td>
<td>32</td>
</tr>
<tr>
<td></td>
<td>1 Type-Y</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

As explained in Section II.D, a limitation of the pilot results is that the smaller group size and payoff incentives, combined with the use of the Strategy Method, created a dominant strategy for unconditional action in the Liability treatment. In the pilot’s decision environment, the initial endowment from not acting was 18 tokens, and the return from acting was 30 tokens. As a result, each Type-X subject was better off acting conditional on at least one other Type-X subject acting (30 – (18/2) > 18). Knowing that the other Type-X subjects had a dominant strategy to act conditional on one other Type-X subject acting, Type-X subjects also had a dominant strategy to act in their unconditional decision, even in the simultaneous decision environment. However, this “limitation” turned out to be a benefit, since we were able to compare two liability regimes where actors had a dominant strategy to act.

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160 With two Type-X subjects equally splitting the 18-token compensation to the Type-Y subject, each Type-X subject is promised a payoff of 21 tokens (30 – 18/2)—a higher payoff compared to the 18 tokens from not acting.

161 To see why, and without limitation, consider XI’s conditional response. If XI’s Action Table indicated that she would not act under any circumstances, the Type-X subject will not act and thus expects a gain of $18. If XI’s Action Table indicated that she would act if one other actor would, the activity will take place only if another Type-X subject acts. In such a case, XI can expect 21 tokens, which is higher compared to not acting (21 > 18). Thus, XI’s best strategy is to act conditional on one other acting. The same is true for all other Type-X subjects. Knowing this, all Type-X subjects would choose to act in their unconditional decision.

162 See supra Section II.D.
While the pilot lacks sufficient power for conducting formal statistical tests, we can examine summary statistics and trends observed in these sessions to see if they are broadly consistent with our hypotheses.\textsuperscript{163}

Table AI-2 reports the proportion of rounds where Type-X subjects chose to act, averaged across all seven rounds. The values in parentheses are the standard deviations of average proportions across Type-X subjects. The larger the standard deviation, the more heterogeneity exists in the data. For instance, the average proportion of rounds where a Type-X subject chose to act unconditionally in CF–L is 91%. While perhaps not surprising given the dominant strategy to act unconditionally, this result indicates that subjects understand and are responding to the incentive scheme. In comparison, subjects chose to act unconditionally in the CF–NL in over 80% of the rounds on average. The fact that subjects chose to act less often in CF–NL is surprising, given that acting when there is No Liability is a clear dominant strategy—in fact, a much clearer dominant strategy than in CF–L. The results may indicate Type-X subjects’ other-regarding preferences,\textsuperscript{164} such as kindness or guilt for the Type-Y subject who was left with zero tokens when one Type-X subject acted. This is also supported by the average unconditional acting of the two treatments in Figure AI-1. The average proportion of subjects acting across all Type-X subjects is lower in all CF–NL rounds compared to CF–L rounds.

\textsuperscript{163} Statistical power is the probability a hypothesis test limits Type II errors or false negative errors. That is, stronger power implies a higher chance of rejecting the null hypothesis when the null hypothesis is in fact false.

\textsuperscript{164} An individual with “other-regarding preferences” is different than one with “self-regarding preferences” in that she also considers other individuals’ utility or consumption levels. For example, someone with altruistic other-regarding preferences may prefer a situation in which others consume or have more wealth, whereas the envious would prefer that others do worse.
Table AI-2: Descriptive Statistics Across All Rounds in Pilot

<table>
<thead>
<tr>
<th></th>
<th>CF–L N = 24</th>
<th>CF–NL N = 24</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Unconditional Decision to Act</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>91.67%</td>
<td>82.74%</td>
</tr>
<tr>
<td></td>
<td>(23.04%)</td>
<td>(37.90%)</td>
</tr>
<tr>
<td><strong>Conditional Decision to Act Upon:</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>0 Other Type-X acting (0X)</td>
<td>12.50%</td>
<td>51.19%</td>
</tr>
<tr>
<td></td>
<td>(30.76%)</td>
<td>(50.14%)</td>
</tr>
<tr>
<td>1 Other Type-X acting (2X)</td>
<td>95.24%</td>
<td>88.10%</td>
</tr>
<tr>
<td></td>
<td>(20.49%)</td>
<td>(32.48%)</td>
</tr>
<tr>
<td>2 Other Type-X acting (3X)</td>
<td>99.40%</td>
<td>93.45%</td>
</tr>
<tr>
<td></td>
<td>(2.92%)</td>
<td>(24.81%)</td>
</tr>
</tbody>
</table>

Figure AI-1: Average Unconditional Decisions in Pilot

Table AI-2 reports average conditional choices to act using the Strategy Method. In CF–L, about 12% of subjects chose to act conditional on no other Type-X subjects acting. It is not clear why a Type-X subject would choose to act in this situation. This result may indicate a small amount of confusion on the part of the subjects while filling out the Action Table. Over 95% of subjects chose to act conditional on one other Type-X subject acting, and over 99% chose to act conditional on two other Type-X subjects acting.

Proportion of Type-X subjects who decided to act, averaged across all seven rounds. Values in parentheses are standard deviations across Type-X subjects’ average decisions to act. The total number of subjects was 48 (24 in the CF–L and 24 in the CF–NL) resulting in 336 observations (48 • 7).
Taken together, these results indicate that subjects’ willingness to act increases with the number of others choosing to act. These findings support our hypotheses and, for the most part, are consistent with the main experiment.

In CF–NL, about 50% of Type-X subjects indicate a willingness to act if they are the only Type-X subject acting. The proportion of Type-X subjects choosing to act surges to 88% or more when there is at least one other Type-X subject acting—a trend similar to what we observed in the main analysis.\textsuperscript{166} This jump occurs despite no change in the monetary incentives across conditional decisions to act in this treatment. This is another instance where it appears that subjects are responding to monetary incentives but not as strongly as one may predict. Subjects may be either self-regarding with a predilection for conforming, or they may have other-regarding preferences, or both.

Compared to CF–L, in CF–NL, more subjects are willing to act conditional on no other Type-X subjects acting—that is, conditional on acting alone. However, conditional on one or two other Type-X subjects acting, when it is a dominant strategy to act under Liability, there is less acting with No Liability. This is illustrated in Figure AI-2, which reports the average proportion of subjects acting in each conditional decision described. These pilot results are consistent with a deterrent effect for liability that switches to an incentive to act once multiple subjects are sharing the costs of providing compensation.

\textsuperscript{166} See supra Table 4.
Figure AI-2: Conditional Decisions to Act Using the Strategy Method in Pilot

**Decision to Act: Conditional-0X-Pilot**

- 3X1Y CF–L
- 3X1Y CF–NL

**Decision to Act: Conditional-1X-Pilot**

- 3X1Y CF–L
- 3X1Y CF–NL

**Decision to Act: Conditional-2X-Pilot**

- 3X1Y CF–L
- 3X1Y CF–NL
The consistency and stability of the pilot results are further supported by the fairness feedback, gathered from subjects’ direct responses to the post-experiment questions, as depicted in Figure AI-3. Subjects were asked to rank on a scale from 0 to 5 the fairness of Type-X subjects acting as a function of the number of actors, with 0 being unfair and 5 being completely fair. The results show that under both CF–NL and CF–L, fairness levels were always above the 2.5 median, that is, the (tortious) activity was considered more fair than unfair. Moreover, under both liability regimes, fairness levels monotonically increased with the number of actors. In other words, subjects in CF–NL seemed to have thought that when three actors inflict an indivisible harm on the Type-Y subject, the result is fairer than two actors doing the same, which is still fairer than one actor solely inflicting the harm. However, compared to the CF–NL, under CF–L, acting was perceived as fairer when two or three Type-X subjects act.

In summary, the pilot results indicate that subjects generally understood the nature of the experiment, instructions, and procedure. The Liability manipulation appears to influence subjects’ behavior in a way that is generally consistent with our design and expectations. Results seem to support our main hypotheses. Subjects were willing to act more often under Liability compared to No Liability, despite having a dominant strategy to act in both treatments, suggesting that shared liability may reduce deterrence (Hypothesis 1). It also appears that subjects were more willing to participate in the tortious activity as the number of actors increased, indicating a tortfest dynamic (Hypothesis 2). Finally, the results support the claim that the but-for test is operable in multiple-causes cases (Hypothesis 3), as subjects were
more willing to act in the tort conditional on others acting—as our theory predicts.

APPENDIX II: DETAILED DESCRIPTION OF THE MAIN RESULTS

This Appendix provides a more detailed description of the statistical analysis reported above. Table AII-1 below is a more comprehensive version of Table 5 in Section II.E.1 and reports more results of our mixed-effect logit regressions. The multilevel model used in these regressions has (1) individual-level random effects to control for correlations between an individual’s decisions over time, as well as (2) session-level random effects to control for correlations between subjects who are randomly rematched within a session by Strangers Matching. Table AII-1 reports an additional variable (not reported in Table 5) called Round. Because each subject participated in seven decision rounds, we test if and in what way a subject’s behavior changed over time. The concern is that subjects might be learning about the incentives of the experiment from one round to another and that this learning effect may drive some of the results. If so, we would expect to see Round significantly affecting the decision to act. However, Round is insignificant in each model, so we can conclude that subjects have a good understanding of the incentives of the experiment in Round 1. Their decisions to act do not change over time.

<table>
<thead>
<tr>
<th>Independent Variables</th>
<th>Model 1 Uncond. Odds Ratio</th>
<th>Model 2 Cond.-0X Odds Ratio</th>
<th>Model 3 Cond.-1X Odds Ratio</th>
<th>Model 4 Cond.-2X Odds Ratio</th>
<th>Model 5 Cond.-3X Odds Ratio</th>
</tr>
</thead>
<tbody>
<tr>
<td>CF–L Dummy</td>
<td>0.25*</td>
<td>0.01***</td>
<td>&lt; 0.01***</td>
<td>1.41</td>
<td>6.01</td>
</tr>
<tr>
<td></td>
<td>(0.19)</td>
<td>(0.01)</td>
<td>( &lt; 0.00)</td>
<td>(1.19)</td>
<td>( &lt; 0.00)</td>
</tr>
<tr>
<td></td>
<td>p = 0.066</td>
<td>p &lt; 0.001</td>
<td>p &lt; 0.001</td>
<td>p = 0.685</td>
<td>p = 0.160</td>
</tr>
<tr>
<td>Round</td>
<td>1.08 (0.08)</td>
<td>1.13 (0.09)</td>
<td>0.08 (0.10)</td>
<td>1.02 (0.10)</td>
<td>0.86 (0.10)</td>
</tr>
<tr>
<td></td>
<td>p = 0.299</td>
<td>p = 0.122</td>
<td>p = 0.387</td>
<td>p = 0.849</td>
<td>p = 0.184</td>
</tr>
</tbody>
</table>

We performed some additional variations in our regression analysis to confirm the robustness of the main results reported in Table 5. Our analysis and results are not qualitatively or statistically altered if we include a lagged

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\[167\] See, e.g., Peter G. Moffatt, Experimetrics: Econometrics for Experimental Economics 92 (2016) (controlling for potential session-effects and correlation of individuals’ contributions in repeated periods); Guillaume R. Fréchette, Session-Effects in the Laboratory, 15 Experimental Econ. 485 (2012).

\[168\] Standard errors in parentheses. Models 1–5: random effects at the individual and session level. \(n = 504\). The omitted dummy variable is CF–NL.

\[169\] See supra Section II.E.1.a.
decision to act as an independent variable, which is another way of testing for dependence in the decisions to act that could influence the results. We also use an alternative estimation strategy: a mixed-effects linear panel regression model. The structure of the logit model we use for our main results is a better fit for the nature of our decision because it corresponds well with the binary choice to act (1) or not (0). However, the linear model is a possible alternative estimation strategy. We use this strategy for Model 3 (1X), because our mixed-effects logit Model 3 did not converge. We continue to control for random effects at the individual level (in case individual specific effects influenced the results). Again, our qualitative and statistical results remained robust.

The analysis in Section II.E.2 looks for statistically significant effects of variation in our liability treatments—Liability versus No Liability—across subjects’ conditional decisions. In order to use a panel-regression approach, individuals can only make one decision in each Round. It is not possible to use a panel-regression approach to simultaneously capture all four conditional decisions in each Round as a common function of the Liability regime and our other controls. Instead, our main results are based on nonparametric tests for conditional decisions averaged across all Rounds. Below, we also report results based on nonparametric tests examining individual decisions from the Round 1 only, as a robustness check. Individual decisions are statistically independent in Round 1 because subjects have not yet received feedback on others’ decisions. These nonparametric tests eliminate one of the dependencies that would otherwise be controlled for in panel regressions.\textsuperscript{170} This observation implies that decisions in Round 1 describe behavior fairly closely to when decisions across all rounds are averaged. The results confirm those reported and discussed in Section II.E.2.\textsuperscript{171}

The analysis in Section II.E.3 involves comparison of results across the Context-Free and Vignette frames. Table AII-2 below reports expanded results for our mixed-effect logit regressions on unconditional decisions to act in these two treatments. Model 6 compares V–L to CF–L, Model 7 compares CF–L to V–NL, and Model 8 compares CF–NL to V–L. With regard to the CF–L treatment: Signed-rank tests, Round 1 shows a statistically significant difference between decisions in Cond-1X and Cond-2X (Cond-0X versus Cond-1X, $p = 0.103$; Cond-1X versus Cond-2X, $p < 0.001$ (this is the same highly significant jump at the point where shared liability makes the tort monetarily attractive); Cond-2X versus Cond-3X, $p = 0.103, n = 36$). With regard to the CF–NL treatment: Signed-rank tests, round 1 shows a statistically significant difference between decisions in Cond-0X and Cond-1X (Cond-0X versus Cond-1X, $p = 0.005$ (this is the highly statistically significant unwillingness to act alone even when there is no liability); Cond-1X versus Cond-2X, $p < 0.103$; Cond-2X versus Cond-3X, $p > 0.999, n = 36$).

\textsuperscript{170} Also, as seen in Figure 3 in Section II.E.2, on average, subjects’ decisions are relatively stable over time.

\textsuperscript{171} With regard to the CF–L treatment: Signed-rank tests, Round 1 shows a statistically significant difference between decisions in Cond-1X and Cond-2X (Cond-0X versus Cond-1X, $p = 0.103$; Cond-1X versus Cond-2X, $p < 0.001$ (this is the same highly significant jump at the point where shared liability makes the tort monetarily attractive); Cond-2X versus Cond-3X, $p = 0.103, n = 36$). With regard to the CF–NL treatment: Signed-rank tests, round 1 shows a statistically significant difference between decisions in Cond-0X and Cond-1X (Cond-0X versus Cond-1X, $p = 0.005$ (this is the highly statistically significant unwillingness to act alone even when there is no liability); Cond-1X versus Cond-2X, $p < 0.103$; Cond-2X versus Cond-3X, $p > 0.999, n = 36$).
compares V–NL to CF–NL, and Model 8 compares V–NL to V–L. Recall that the Round dummy variable captures any change in willingness to act as individuals complete subsequent rounds of the experiment. The Round dummy variable is greater than 1 in Models 6 through 8 and significant in Models 7 and 8. This indicates that subjects are becoming more willing to act unconditionally as the experiment progresses. It appears this is particularly true for V–NL (as Round is insignificant in Model 6 with V–L).

<table>
<thead>
<tr>
<th>Independent Variables</th>
<th>Model 6</th>
<th>Model 7</th>
<th>Model 8</th>
</tr>
</thead>
<tbody>
<tr>
<td>V–L Dummy</td>
<td>8.78 (7.19)</td>
<td>—</td>
<td>0.65 (0.38)</td>
</tr>
<tr>
<td></td>
<td>p = 0.008</td>
<td>—</td>
<td>p = 0.453</td>
</tr>
<tr>
<td>V–NL Dummy</td>
<td>—</td>
<td>2.54 (1.33)</td>
<td>—</td>
</tr>
<tr>
<td></td>
<td>—</td>
<td>p = 0.075</td>
<td>—</td>
</tr>
<tr>
<td>Round</td>
<td>1.08 (0.09)</td>
<td>1.17 (0.09)</td>
<td>1.20 (0.10)</td>
</tr>
<tr>
<td></td>
<td>p = 0.306</td>
<td>p = 0.030</td>
<td>p = 0.027</td>
</tr>
</tbody>
</table>

In Section II.E.3, we also discuss the Vignette treatment’s impact on conditional willingness to act—that is, we compare V–L to CF–L. Table AII-3 provides expanded results for Models 9 through 12 (the four conditional decisions) in Table 7. These results are the analog to those for the Context-Free frame in Table 5 and Table AII-1. The Round variable is not significant in any Model, implying subjects’ conditional decisions are relatively stable over decision rounds.

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172 Standard errors in parentheses. Random effects at the individual and session level.
Model 6: sessions = 6, subjects = 76, n = 532. Omitted dummy variable: CF–L.
Model 7: sessions = 6, subjects = 84, n = 588. Omitted dummy variable: CF–NL.
Model 8: sessions = 6, subjects = 88, n = 616. Omitted dummy variable: V–NL.
Table AII-3: Mixed-Effects Logit Regression—Liability Contextual Treatment Comparisons

<table>
<thead>
<tr>
<th>Independent Variables</th>
<th>Model 9 Cond.-0X Odds Ratio</th>
<th>Model 10 Cond.-1X Odds Ratio</th>
<th>Model 11 Cond.-2X Odds Ratio</th>
<th>Model 12 Cond.-3X Odds Ratio</th>
</tr>
</thead>
<tbody>
<tr>
<td>V–L Dummy</td>
<td>0.81 (0.91)</td>
<td>0.52 (0.71)</td>
<td>0.60 (0.50)</td>
<td>0.74 (0.96)</td>
</tr>
<tr>
<td></td>
<td>( p = 0.850 )</td>
<td>( p = 0.632 )</td>
<td>( p = 0.547 )</td>
<td>( p = 0.815 )</td>
</tr>
<tr>
<td>Round</td>
<td>0.91 (0.10)</td>
<td>1.18 (0.13)</td>
<td>1.05 (0.10)</td>
<td>0.85 (0.12)</td>
</tr>
<tr>
<td></td>
<td>( p = 0.378 )</td>
<td>( p = 0.140 )</td>
<td>( p = 0.636 )</td>
<td>( p = 0.269 )</td>
</tr>
</tbody>
</table>

As noted above, it is not possible to use a panel data approach to simultaneously capture the impact of varying the Vignette–Liability treatment across all four conditional decisions in each Round. We generate results for this discussion in Section II.E.3 by using the same nonparametric approach discussed above for the Context-Free frame. Critical robustness checks include: V–L willingness to act is significantly higher conditional on 2X than 1X acting (Wilcoxon signed-rank test using only Round 1 data, \( p < 0.001 \), \( n = 40 \)); V–NL willingness to act is significantly higher conditional on 1X acting than 0X (Wilcoxon signed-rank test using only Round 1 data, \( p < 0.001 \), \( n = 48 \)).

Section II.E.4 discusses our fairness-rating results, both across liability treatments and as the number of Type-X acting increased. We use individual-level fairness rankings to conduct our nonparametric statistical analysis. Recall that our experimental procedure used Strangers Matching to randomly rematch individuals in each round. The most conservative approach to statistically test for fairness ratings would have been to focus on session-level measures. This approach would fully control for statistical dependency between subjects within a session but would generate only six observations per liability treatment. We believe our less conservative approach is reasonable with Strangers Matching, but as we do not have a full robustness check for this type of dependency, this remains a caveat to our fairness rating results. Such a caveat is not required for the primary analysis of the decision to act.

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173 Standard errors in parentheses. Models 9–12: random effects at the individual and session level. \( n = 532 \). Omitted dummy variable is CF–L.

174 Robustness checks on the analysis of the decision to act are possible because subjects made decisions to act across seven rounds. Thus, where possible, we also used panel regressions examining the decisions to act that control for individual-level and session-level dependencies as robustness checks for the individual-level nonparametric tests. In instances where panel regressions were not possible, nonparametric tests were conducted using only Round 1 data, as subjects were statistically independent at that point in the experiment.
In addition to the core individual-choice and fairness-perception data that we analyze in the paper, we also collected information about a number of additional variables for our participants. Table AII-4 above reports the summary statistics for the variables used as additional controls in the regression analysis discussed in Section II.E.5. These additional controls generally had no statistically significant impact on our results. To illustrate, Table AII-5 provides results for unconditional decisions—Vignette and Context-Free combined—using the multilevel logit regression analysis with our additional controls. Results from additional regressions and analysis integrating these controls were also generally insignificant and revealed no systematic impact from these controls on our results.

The summary statistic for Economic Experiments is the average number of other economics experiments subjects have previously participated in. The summary statistic for Risk Intensity is average risk intensity on a scale from 0 to 10, with 0 being extremely risk-seeking and 10 being extremely risk averse. The summary statistics for dummies are the percentages of subjects who responded affirmatively to those categories.

Due to the small number of observations, transgender subjects were dropped for this analysis (three subjects, twenty-one observations). Statistical inferences from only three subjects are problematic.
**Table AII-5: Panel Logit Regression with Additional Risk and Demographic Controls**

<table>
<thead>
<tr>
<th>Independent Variables</th>
<th>Uncond. No Liability Odds Ratios</th>
<th>Uncond. Liability Odds Ratio</th>
</tr>
</thead>
<tbody>
<tr>
<td>Vignette Dummy</td>
<td>3.51** (1.90) $p = 0.020$</td>
<td>6.28** (5.44) $p = 0.034$</td>
</tr>
<tr>
<td># Other Economic Experiments</td>
<td>1.02 (0.19) $p = 0.919$</td>
<td>0.73 (0.17) $p = 0.167$</td>
</tr>
<tr>
<td>Risk Intensity (0–10)</td>
<td>1.16 (0.17) $p = 0.295$</td>
<td>0.98 (0.26) $p = 0.943$</td>
</tr>
<tr>
<td>Male Dummy</td>
<td>2.45 (1.46) $p = 0.130$</td>
<td>0.33 (0.30) $p = 0.228$</td>
</tr>
<tr>
<td>Business-Major Dummy</td>
<td>0.42 (0.25) $p = 0.149$</td>
<td>3.95 (4.01) $p = 0.176$</td>
</tr>
<tr>
<td>White Dummy</td>
<td>0.55 (0.37) $p = 0.380$</td>
<td>4.61 (4.51) $p = 0.119$</td>
</tr>
<tr>
<td>Believers Dummy</td>
<td>1.08 (0.57) $p = 0.886$</td>
<td>4.22 (3.90) $p = 0.118$</td>
</tr>
<tr>
<td>Democrat Dummy</td>
<td>0.55 (0.33) $p = 0.321$</td>
<td>0.83 (0.89) $p = 0.862$</td>
</tr>
<tr>
<td>Republican Dummy</td>
<td>3.14 (2.40) $p = 0.135$</td>
<td>0.38 (0.43) $p = 0.396$</td>
</tr>
<tr>
<td>Political-Party Intensity (0–5)</td>
<td>1.04 (0.22) $p = 0.859$</td>
<td>1.06 (0.33) $p = 0.857$</td>
</tr>
<tr>
<td>Round</td>
<td>1.21 (0.09) $p = 0.016$</td>
<td>1.08 (0.09) $p = 0.374$</td>
</tr>
<tr>
<td>Characteristic</td>
<td>Out of 200 subjects</td>
<td></td>
</tr>
<tr>
<td>--------------------------------------</td>
<td>---------------------</td>
<td></td>
</tr>
<tr>
<td>Previous Economics Experiments</td>
<td>1.215</td>
<td></td>
</tr>
<tr>
<td>Risk Intensity (0–10)</td>
<td>4.545</td>
<td></td>
</tr>
<tr>
<td>Males</td>
<td>84 (3 transgender/other)</td>
<td></td>
</tr>
<tr>
<td>Business Major</td>
<td>85</td>
<td></td>
</tr>
<tr>
<td>White</td>
<td>157</td>
<td></td>
</tr>
<tr>
<td>Believers</td>
<td>115</td>
<td></td>
</tr>
<tr>
<td>Democrat</td>
<td>59</td>
<td></td>
</tr>
<tr>
<td>Republican</td>
<td>42</td>
<td></td>
</tr>
<tr>
<td>Political-Party Intensity (0–5)</td>
<td>2.33</td>
<td></td>
</tr>
<tr>
<td>Age</td>
<td>22.025</td>
<td></td>
</tr>
</tbody>
</table>